Appendix: What’s Different about

Bank Holding Companies?

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March 9, 2018

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1 Trading Desk’s Problem

The trading desk’s problem (30) subject to (31) is solved by specifying the Hamilton-Jacobi-Bellman (HJB) equation when the change in the lifetime utility is found using (30) and (3) in the paper. After choosing the optimal portfolio the manager’s problem boils down to solving a Partial Differential Equation (PDE) for the lifetime utility \( h(\tau, X) \). Here \( h(\tau, X) \) is \( J(\tau, X) \) in equation (32) of the paper. This lifetime utility under the optimal behavior must be the solution to the PDE.

\[
\frac{\partial h(\tau, X)}{\partial \tau} = \frac{1}{2} \text{Trace} \left( A' \frac{\partial^2 h(\tau, X)}{\partial X \partial X} A \right) - 2 \left( \frac{\partial h(\tau, X)}{\partial X} \right)' \left[ B X + C \right] - G \\
- h(\tau, X) \left[ X'DX + EX + F \right] - \frac{1}{2h(\tau, X)} \text{Trace} \left( H \frac{\partial h(\tau, X)}{\partial X} \left( \frac{\partial h}{\partial X} \right)' \right)
\]

subject to

\[
h(0, X) = h(X).
\]

\( h(X) \) is some given terminal lifetime utility of the investor.

The coefficients are given by

\[
A \equiv \Sigma_X \\
B \equiv \frac{1}{2} \left[ A^P - (\gamma^j - 1) \Sigma_X \Sigma_X' (b' - \iota b_n) \omega_1 (b - \iota b_n) (A^P - A^Q) \right] \\
C \equiv \frac{1}{2} \left[ -\xi \Sigma_X \Sigma_X' b_n + (\gamma^j - 1) \Sigma_X \Sigma_X' (b - \iota b_n)' \omega_1 K - \gamma^P \right] \\
D \equiv \frac{\gamma^j - 1}{2 \gamma^j} (A^P - A^Q)' (b' - \iota b_n') \omega_1 (b - \iota b_n') (A^P - A^Q) \\
E \equiv \frac{\gamma^j - 1}{\gamma^j} (\delta_1 - \xi b_n (A^P - A^Q)) - \frac{\gamma^j - 1}{\gamma^j} K' \omega_1 (b - \iota b_n) (A^P - A^Q) \\
F \equiv \frac{1 - \gamma^j}{2} \xi^2 b_n \Sigma_X \Sigma_X' b_n' + \frac{\gamma^j - 1}{2 \gamma^j} K' \omega_1 K + \frac{\beta}{\gamma^j} + \frac{\gamma^j - 1}{\gamma^j} \left[ \delta_0 + \xi b_n (\gamma^P - \gamma^Q) \right]
\]

with

\[
K \equiv (b - \iota b_n) (\gamma^P - \gamma^Q) - \gamma^j (b \Sigma_X \Sigma_X' b_n' - \iota b_n \Sigma_X \Sigma_X' b_n') \xi
\]

\[
G \equiv -\beta \gamma^j
\]

\[
H \equiv (\gamma^j - 1) \left[ \gamma^j \Sigma_X \Sigma_X' (b' - \iota b_n') \omega_1 (b - \iota b_n) - I_n \right] \Sigma_X \Sigma_X'
\]

In the text we use four treasury securities so that \( b_{4\tau} \) is used for the generic \( b_n \).

\( G = 0 \) when the trading desk does not consider periodic withdrawals from the portfolio.
Table 1: Estimates of Parameters for PDE (1).

<table>
<thead>
<tr>
<th>A</th>
<th>B</th>
<th>C</th>
<th>D</th>
<th>E</th>
<th>F</th>
<th>H</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.0313</td>
<td>0.2716</td>
<td>0.0119</td>
<td>57.5221</td>
<td>3.2240</td>
<td>0.0828</td>
<td>0</td>
</tr>
</tbody>
</table>

The coefficients in the PDE (1) are in Table 1 for the parameters from the term structure model and the preference parameters $\gamma^j = 10, \beta = 0.05, \xi = 1.0$.

Sangvinatsos and Wachter (2005, p. 192 JF) guess the solution when the trading desk does not make periodic withdrawals between meetings of the COO and trading desk.

$$h(\tau, X) = \exp \left\{ -\frac{1}{2} X' B_3(\tau) X + B_2(\tau)' X + B_1(\tau) \right\},$$

where $\tau = T - t$.

This may be written as

$$h(\tau, X) = \exp \left\{ -\frac{1}{2} (X - (B_3(\tau))^{-1} B_2(\tau)) B_3(\tau) (X - (B_3(\tau))^{-1} B_2(\tau)) \right.$$  

$$+ B_1(\tau) + \frac{1}{2} B_2(\tau)' (B_3(\tau))^{-1} B_2(\tau) \right\} \tag{5}$$

so that $(B_3(\tau))^{-1} B_2(\tau)$ is the mean and $B_3(\tau)^{-1}$ is the variance of the expected utility of terminal wealth. We call these terms $\mu_j(\tau)$ and $\sigma_j(\tau)$ in equation (32) of the paper.

If one takes the derivatives of the guess (4) and substitute into the linear PDE, then one gets the Ricatti ordinary differential equations. The quadratic form matrix satisfies the ODE

$$\frac{\partial B_3(\tau)}{\partial \tau} = -B_3(\tau) A[I - H] A' B_3(\tau) - 2B_3(\tau) B + D \tag{6}$$

subject to

$$B_3(0) = B_3.$$

The first line uses the symmetry of $A$ so that $A' = A$. In addition, the matrix $B_3(\tau)$ must be positive definite, which is true when $A[I - H] A'$ and $D$ are positive definite.

The ODE for the linear coefficients is

$$\frac{\partial B_2(\tau)}{\partial \tau} = -B_2(\tau)' A[I - H] A' B_3(\tau) + C' B_3(\tau) - B_2(\tau)' B - E \tag{7}$$
subject to

\[ B_2(0) = B_2. \]

The ODE for the constant coefficients yields

\[
B_1(\tau) = B_1(0) + \frac{1}{2} \int_0^\tau \left[ B_2(s)' A [I - H] A' B_2(s) - \text{Trace} (AA'B_3(s)) - 2B_2(s)'C - 2F \right] ds
\]

with

\[ B_1(0) = B_1. \]

These expressions are similar to Sangvinatsos and Wachter (2005, p. 222 and 223). The solutions to these ODEs under the affine term structure estimates are given in Table 2. The results are reported in Table 2 for \( \gamma = 10 \) and \( \beta = 0.05 \) at a one year time horizon. Figure 1 provides a graph for this solution versus the level of the yield curve in the left hand graph. The domain is ±3 standard deviation of the lifetime utility relative to its mean. The right hand graph has the hedging demand for 5 year bonds, and the total demand for 3 Months and 5 Year government bonds following (8).

<table>
<thead>
<tr>
<th>Table 2: Solution to the ODEs (6), (7), and (8)</th>
</tr>
</thead>
<tbody>
<tr>
<td>( B_1(1) )</td>
</tr>
<tr>
<td>-0.0770</td>
</tr>
</tbody>
</table>

Figure 1: The Expected Lifetime Utility of the trading desk and Portfolio Weights

Table 4 in the paper comes from Table 2 and Figure 6 in the paper corresponds to 1.
We want to find the stochastic process for the lifetime utility given the solution to the PDE (1) for $h(X,t)$ and the optimal portfolio rule (33) in the paper.

First we find the stochastic process for the trading desk’s utility from her bank capital, since the lifetime utility follows (32) in the paper.

\[ J(\tau, K_M^j, X) = \frac{1}{1-\gamma^j} \left( K_M^j \right)^{1-\gamma^j} h^{\gamma^j}(\tau, X) \quad (9) \]

The stochastic process for the trading desk’s utility from capital is

\[
\frac{d \left( K_M^j \right)^{1-\gamma^j}}{(K_M^j)^{1-\gamma^j}} = \left( C_1(\tau) - \frac{1}{2} \left( X(s)'C_3(\tau)X(s) - 2C_2(\tau)X(s) \right) \right) ds + \left( C_4(\tau) + X(s)'C_5(\tau) \right) d\epsilon_s. \quad (10)
\]

\[
C_1(\tau) \equiv (1-\gamma^j) \left\{ \delta_0 + \xi b_n'(\gamma^P - \gamma^Q) - \frac{1}{2} \gamma^j \xi^2 b_{4r} \Sigma_X \Sigma_X' X_n b_{4r} + K' \omega_1 K_w + \frac{1}{2} \gamma^j K' \omega_1 K \\
+ \gamma^j (K'_w - K') \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_2(\tau) \\
- \frac{1}{2} (\gamma^j)^2 B_2(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_2(\tau) \right\},
\]

\[
C_2(\tau) \equiv (1-\gamma^j) \left[ \delta_1 - \xi b_n'(A^P - A^Q) - (K'_w + (1-\gamma^j) K') \omega_1 (b - \nu b_n)(A^P - A^Q) \\
+ \gamma^j (K'_w - K) \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_3(\tau) \\
+ \gamma^j B_2(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \omega_1 (b - \nu b_n) (A^P - A^Q) \\
- (\gamma^j)^2 B_2(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_3(\tau) \right], \quad (11)
\]

\[
C_3(\tau) \equiv (1-\gamma^j) \left[ \gamma^j (A^P - A^Q)' (b' - \nu b_n) \omega_1 (b - \nu b_n) (A^P - A^Q) \\
- 2\gamma^j (A^P - A^Q)' (b' - \nu b_n) \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_3(\tau) \\
+ (\gamma^j)^2 B_3(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \omega_1 (b - \nu b_n) \Sigma_X \Sigma_X' B_3(\tau) \right],
\]

\[
C_4(\tau) \equiv (1-\gamma^j) \left[ K + \gamma^j B_2(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \right] \omega_1 (b - \nu b_n) \Sigma_X + (1-\gamma^j) \xi b_n \Sigma_X \\
C_5(\tau) \equiv (1-\gamma^j) \left[ -(A^P - A^Q)' (b' - \nu b_n) + \gamma^j B_3(\tau)' \Sigma_X \Sigma_X' (b' - \nu b_n) \right] \omega_1 (b - \nu b_n) \Sigma_X,
\]

\[
K_w \equiv (b' - \nu b_n)(\gamma^P - \gamma^Q) - \gamma^j (b \Sigma_X \Sigma_X' b_n - \nu b_n \Sigma_X \Sigma_X' b_n)' \xi.
\]
Table 3 gives the values of the coefficients $C_i(\tau)$ for $\tau = 1$ given by (11). In this case the discounted future wealth is positively related to the future factor for low values of the level factor, i.e. $X_s < - \left( C_3(\tau) \right)^{-1} C_2(\tau) < 0$.

**Table 3: Estimates of Parameters for Equation (11) with $\gamma^j = 0$.**

<table>
<thead>
<tr>
<th>$C_1(\tau)$</th>
<th>$C_2(\tau)$</th>
<th>$C_3(\tau)$</th>
<th>$C_4(\tau)$</th>
<th>$C_5(\tau)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.0725</td>
<td>-11.5339</td>
<td>244.3535</td>
<td>-0.1512</td>
<td>-2.2074</td>
</tr>
</tbody>
</table>

By using Ito’s Lemma again one finds:

$$J(w, X, \tau) = J(w, X, 0) \exp \left\{ \int_0^{\tau} \left( J_1(0) - \frac{1}{2} \left( X(s)'J_3(0)X(s) - 2J_2(0) \right) \right) ds \right. 
- \left. \int_0^{\tau} \left( J_4(0) + X(s)'J_5(0) \right) d\epsilon_s \right\}. \quad (12)$$

The coefficients

$$J_1(0) \equiv C_1(0) - \frac{\gamma^j}{2} (1 - \gamma^j) C_4(0) C_4(0)' + \gamma^j F + \frac{\gamma^j}{2} B_2(\tau)' \left[ (\gamma^j + 1) \Sigma X \Sigma X' + H \right] B_2(\tau)$$
- $\gamma^j B_2(\tau)' \left( (1 - \gamma^j) \Sigma X C_4(0)' + 2C - \gamma P \right)$

$$J_2(0) \equiv C_2(0) - \gamma^j (1 - \gamma^j) C_4(0) C_5(0)' + \gamma^j E + \gamma^j B_2(\tau)' \left[ (\gamma^j + 1) \Sigma X \Sigma X' + H \right] B_3(\tau)$$
- $\gamma^j B_2(\tau)' \left( (1 - \gamma^j) \Sigma X C_5(0)' - 2B + A P \right)$

$$J_3(0) \equiv C_3(0) + \gamma^j (1 - \gamma^j) C_5(0) C_5(0)' - 2\gamma^j D + \gamma^j B_3(\tau)' \left[ (\gamma^j + 1) \Sigma X \Sigma X' + H \right] B_3(\tau)$$
- $2\gamma^j B_3(\tau)' \left( (1 - \gamma^j) \Sigma X C_5(0)' - 2B + A P \right)$,

$$J_4(0) \equiv C_4(0) + \gamma^j B_2(\tau)' , \text{ and}$$

$$J_5(0) \equiv C_5(0) + \gamma^j B_3(\tau)' .$$

Recall that the coefficients $A, \cdots, H$ are defined in equation (3).

We want $J(w, X, \tau)$ to be a uniformly integrable martingale. We recognize that it is a stochastic exponential (Doléans-Dade exponential). See Protter (pp. 84-89). In our case, we have a continuous stochastic process for the factor. As a result, we have

$$\mathcal{E}(X_t) = exp \left\{ X_t - \frac{1}{2} [X, X]_t \right\} ,$$

where $[X, X]_t$ is the quadratic variation of $J(w, X, \tau)$. 

5
Theorem 45 of Protter (2005, p.141) demonstrates $J(w, X, \tau)$ to be a uniformly integrable martingale as long as

$$E\left[exp\left\{\frac{1}{2}[X, X]_t\right\}\right] < \infty.$$ 

In this case, the quadratic variation includes all the terms associated with the variance-covariance matrix $\Sigma_X \Sigma'_X$. In this case the quadratic variation is

$$E\left\{\exp\left[\left(J_4(0) + X(s)'J_5(0)\right)\left(J_4(0) + X(s)'J_5(0)\right)\right]\right\} < \infty. \quad (14)$$

This is called the Novikov’s Criterion. Cosimano and Jun (2018) show these expectations are bounded for the investor’s problem.

If this is true, then the lifetime utility of the investor is given by

$$J(w, X, \tau) = J(w, X, 0)E_t\left\{\exp\left\{\int_0^\tau \left(J_1(0) - \frac{1}{2}\left(X(s)'J_3(0)X(s) - 2J_2(0)\right)\right) ds\right\}\right\}. \quad (15)$$

To solve the COO’s problem (42) in the paper, we need to find the probability distribution of expressions like (15). It turns out that (15) has the same functional form as the stochastic discount factor. The difference is that the coefficients are different for each stochastic process.

2 Gross Growth Rate of the Trading Desk’s Capital

In Table 4 we apply the forward Kolmogorov equation to the gross rate of growth of the trading desk’s capital using the coefficients $C_i(\tau)$ for $i = 1, 2, 3, 4, 5$ in (11) for $\gamma^j = 0$. This corresponds to equation (36) in the paper. We can also express in Table 5 the conditional probability of this growth rate of capital as in the case of the stochastic discount factor (11) in the paper. This leads to equation (35) in the paper. Tables 4 and 5 are used to construct Table 5 in the paper.

The expected gross growth rate of the capital for the trading desk is a Gaussian distribution in Figure 8 in the paper. At the stationary point for the level of the yield curve $-0.177$ we have $E_t\left[\frac{K_{t+\tau}}{K_{t,t}}\right] = 1.1048$. With the highest level of the yield curve at $0.0256$ we have $E_t\left[\frac{K_{t+\tau}}{K_{t,t}}\right] = 0.8529$, since the higher level leads to a decrease in the price of bonds. In

\footnote{See 24 in Cosimano and Jun (2018).}
Table 4: Solution to Forward Kolmogorov Equation for coefficients in the Gross Growth Rate of the Trading Desk’s Capital.

<table>
<thead>
<tr>
<th>$K_1$</th>
<th>$K_2$</th>
<th>$K_3$</th>
<th>$\sigma_K$</th>
</tr>
</thead>
<tbody>
<tr>
<td>-0.0130</td>
<td>0.0000</td>
<td>1.0895 $10^{-4}$</td>
<td>0.0104</td>
</tr>
</tbody>
</table>

Table 5: Coefficients for Conditional Expectation for the Trading Desk’s Capital.

<table>
<thead>
<tr>
<th>$\mathcal{K}_1$</th>
<th>$\mathcal{K}_2$</th>
<th>$\mathcal{K}_3$</th>
<th>$\mathcal{K}_3^{-1}\mathcal{K}_2$</th>
<th>$\exp\left{-\frac{1}{2}\mathcal{K}_2\mathcal{K}_3^{-1}\mathcal{K}_2 + \mathcal{K}_1 - \frac{1}{2}\mathcal{K}_1(\tau)\right}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.0074</td>
<td>-5.6309</td>
<td>88.1733</td>
<td>-0.0639</td>
<td>1.2138</td>
</tr>
</tbody>
</table>

addition, a higher level should revert to the lower mean, resulting in an expected loss on the portfolio.

We will also deal with the product of the pricing kernel with the gross growth rate of the trading desk’s capital. Its coefficients are given by

\[
D_1(\tau) \equiv C_1(\tau) + \mathcal{M}_1, \\
D_2(\tau) \equiv C_2(\tau) + \mathcal{M}_2, \\
D_3(\tau) \equiv C_3(\tau) + \mathcal{M}_3, \\
D_4(\tau) \equiv C_4(\tau) + \mathcal{M}_4, \quad \text{and} \\
D_5(\tau) \equiv C_5(\tau) + \mathcal{M}_5. 
\tag{16}
\]

Table 6: Estimates of Parameters for Equation (16).

<table>
<thead>
<tr>
<th>$D_1(\tau)$</th>
<th>$D_2(\tau)$</th>
<th>$D_3(\tau)$</th>
<th>$D_4(\tau)$</th>
<th>$D_5(\tau)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.3937</td>
<td>-38.3567</td>
<td>1,522.60</td>
<td>0.5711</td>
<td>33.5455</td>
</tr>
</tbody>
</table>

Tables 7 and 8 apply the forward Kolmogorov equation to find the probability distributions for this product and its conditional expectation which corresponds to equation 38 in the paper.

At the stationary point for the level of the yield curve $-0.177$ we have $E_t\left[\frac{MK_{\tau,t}}{MK_{t,t}}\right] = 1.0917$. With the highest level of the yield curve observed at 0.0256 we have $E_t\left[\frac{MK_{\tau,t}}{MK_{t,t}}\right] = 0.4935$. Finally, the maximum value is $E_t\left[\frac{MK_{\tau,t}}{MK_{t,t}}\right] = 1.1339$.  


Table 7: Solution to Forward Kolmogorov Equation for coefficients in (16).

<table>
<thead>
<tr>
<th>KM_1</th>
<th>KM_2</th>
<th>KM_3</th>
<th>( \sigma_{KM} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>-0.0825</td>
<td>0.0000</td>
<td>( 1.1222 \times 10^{-4} )</td>
<td>0.0106</td>
</tr>
</tbody>
</table>

Table 8: Coefficients for the Conditional Expectation under the Product of Pricing Kernel with Trading Desk’s Capital.

<table>
<thead>
<tr>
<th>MK_1</th>
<th>MK_2</th>
<th>MK_3</th>
<th>( \exp \left{ -\frac{1}{2} MK_1 MK_3^{-1} MK_2 + MK_1 - \frac{1}{2} MK_1(\tau) \right} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>-0.1536</td>
<td>-16.1705</td>
<td>549.4276</td>
<td>-0.0294</td>
</tr>
</tbody>
</table>

Therefore we have expressed all the stochastic processes for solving the loan desk’s problem as standard normal random variables with mean zero. This can be done by adding

\[
\exp \left\{ -\frac{1}{2} A_1(\tau, X) \right\}
\]

to the constant term in \( MK(\tau, X) \).

We write these probabilities as

\[
\frac{1}{M(\tau, X)} \frac{M_{t,t}}{M_{t,t}} = p_M(t, X, \tau, Y), \quad \frac{1}{K(\tau, X)} \frac{K^j_{M}(t+\tau)}{K^j_{M}(t)} = p_K(t, X, \tau, Y), \quad (17)
\]

\[
\frac{1}{KM(\tau, X)} \frac{M_{t,t} K^j_{M}(t+\tau)}{K^j_{M}(t)} = p_{MK}(t, X, \tau, Y), \quad \frac{1}{MP(\tau, X)} \frac{M_{t,t} P_{3t}}{P_{3t}} = p_{MP}(t, X, \tau, Y),
\]

\[
\frac{1}{MY(\tau, X)} \frac{M_{t,t} Y}{M_{t,t}} = p_{MY}(t, X, \tau, Y) \quad \text{and} \quad \frac{1}{KM\bar{Y}(\tau, X)} \frac{M_{t,t} K^j_{M}(t+i\tau)}{K^j_{M}(t)} = p_{MKY}(t, X, \tau, Y).
\]

3 Loan Desk’s Optimization Problem

This section derive additional results for the model in section 5 of the paper. The Lagrangian function for the COO’s problem after using the balance sheet constraint for the loan desk to
remove the quantity of deposits is given by

\[
V(t, K^j_M(t), K^j_L(t), r^j_{2,t}, X(t)) = K^j_M(t)\mathcal{M}(\tau, X)\mathcal{K}(\tau, X) - \frac{r^p}{\bar{D}}\mathcal{M}(\tau, X) [\bar{D} - K^j_M(t)]^2
\]

\[
+ \max_{\tau} \mathcal{E}_t \left\{ \mathcal{M}(\tau, X) \left[ (r^j_{\tau, t} - c^j - r^D_{\tau, t}) \left\{ \gamma^j_{0,0} - \gamma^j_{1,0} r^j_{\tau, t} + \sigma (r^j_{\tau, t} - \bar{r}^j_{\tau, t}) \right\} 
+ (r^j_{2,t} - c^j - r^D_{\tau, t}) \left\{ \gamma^j_{0,0} - \gamma^j_{1,2} r^j_{2,t} + \sigma (r^j_{2,t} - \bar{r}^j_{2,t}) \right\} + r^D_{\tau, t} (K^j_L(t) + (1 - \xi)K^j_M(t) - R_{t}^j) 
+ (r^j_{2,t} - c^j - r^D_{\tau, t}) \left\{ \gamma^j_{0,2} - \gamma^j_{1,2} r^j_{2,t} + \sigma (r^j_{2,t} - \bar{r}^j_{2,t}) \right\} - (1 - \chi)\bar{V}^j_{\tau, t} + (1 - \eta)q^j \right\} \tau 
+ \lambda_1 \mathcal{M}(\tau, X) \left[ K^j_M(t) + K^j_L(t) - \kappa_L \left( L^j_{\tau, t} + L^j_{2,t} + L^j_{2,t} - \gamma_{2,t} \right) - \kappa_T \xi K^j_M(t) - \bar{c}_b (\frac{P_{3r,t}}{P_{3r,t}} - 1) \right] 
+ \lambda_2 \mathcal{M}(\tau, X) \left[ K^j_L(t) - \alpha - \gamma_{2,t} (L^j_{\tau, t} + L^j_{2,t} - \gamma_{2,t}) - \alpha_T \xi K^j_M(t) + \alpha_R R_{t}^j \right] 
+ \mathcal{E}_t \left[ \frac{M_{2r,t}}{M_{t,t}} V(t + \tau, K^j_M(t + \tau), K^j_L(t) + \left[ \pi^j_L(t) - \pi^j_K(t) + q^j \right] \tau, r^j_{2,t}, X(t + \tau) \right] \right\} \right).
\]

(18)

For the two period loans we have

\[
\mathcal{M}(\tau, X) \left[ 2 (-\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t}) r^j_{2,t} - (c^j + r^D_{2,t}) \left\{ -\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t} \right\} + \gamma^j_{0,0} - \gamma^j_{1,2} r^j_{2,t} + \sigma (r^j_{2,t} - \bar{r}^j_{2,t}) \tau - \mathcal{M}(\tau, X) \left[ \lambda_1 \kappa_L + \lambda_2 \alpha_2 \right] \left( -\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t} \right) \right] 
+ \frac{M_{2r,t}}{M_{t,t}} \frac{\partial \mathcal{V}}{\partial m^j_{2,t}} \right] + \mathcal{E}_t \left[ \frac{M_{2r,t}}{M_{t,t}} \frac{\partial \mathcal{V}}{\partial m^j_{2,t}} \right] \frac{\partial \pi^j_L}{\partial m^j_{2,t}} = 0.
\]

Here,

\[
\frac{\partial \pi^j_L}{\partial m^j_{2,t}} = 2 \left\{ -\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t} \right\} \left[ 2 r^j_{2,t} - (c^j + r^D_{2,t}) \left\{ -\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t} \right\} \right],
\]

and

\[
\frac{\partial \mathcal{V}}{\partial m^j_{2,t}} = \frac{\partial \pi^j_L}{\partial m^j_{2,t}} - \left[ \lambda_1 \kappa_L + \lambda_2 \alpha_2 \right] \left( -\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t} \right).
\]

Using the optimal condition for paying dividends (44) in the paper, the optimal condition for the two period loan margin becomes

\[
\left[ 2 r^j_{2,t} - (c^j + r^D_{2,t}) + \frac{\gamma^j_{0,0} + \sigma_0 \bar{r}^j_{2,t}}{\gamma^j_{1,2} + \sigma_1 \bar{r}^j_{2,t}} \right] \tau = \left[ \mathcal{M}(\tau, X) + \mathcal{M}(2\tau, X) \right] \left( \lambda_1 \kappa_L + \lambda_2 \alpha_2 \right).
\]

(19)
Here, the stochastic discount factor is given by (10) in the paper.  
We can also examine the behavior of the two period loans. We focus on the capital constraint (21) in the paper. First we solve (19), when these constraints do not bind, then

$$r^*_{2r,t} = \frac{1}{2} \left( c^j + r^D_{2r,t} \right) - \frac{\gamma^j_0_{2r} + \sigma_0 \varepsilon^j_{2r,t}}{2 \left( -\gamma^j_{1,2r} + \sigma_1 \varepsilon^j_{2r,t} \right)}.$$  

Consequently, the two period loan follows the same rule as one period loans using the demand for two period loans.

When the liquidity and capital constraints are binding, the analysis follows the same argument as for the one period loan in the paper. First the loan rate is set on the demand for two period loans.

$$r^K_{2r,t} = \frac{1}{\left( -\gamma^j_{1,2r} + \sigma_1 \varepsilon^j_{2r,t} \right)} \left[ L^j_{2r,t} - \left( \gamma^j_0_{2r} + \sigma_0 \varepsilon^j_{2r,t} \right) \right].$$  

The subscript \( '2r, t' \) in the two period loan rate refers to the loan rate when the two period loans just satisfy the capital constraint (21) at time \( t \) in the paper.

The lagrange multiplier for the capital constraint (21) at time \( t \) in the paper is found by solving (19) for \( \lambda_1 \) when \( \lambda_2 = 0 \). In addition, we use (21).

$$\lambda_1 = \frac{2 \tau \left[ \chi M(\tau, X) + M(2\tau, X) \right]}{\kappa L \left[ M(\tau, X) + M(2\tau, X) \right]} \left( r^K_{2r,t} - \frac{1}{2} (c^j + r^D_{2r,t}) - \frac{\gamma^j_0_{2r} + \sigma_0 \varepsilon^j_{2r,t}}{2 \left( -\gamma^j_{1,2r} + \sigma_1 \varepsilon^j_{2r,t} \right)} \right).$$

By using (20), the lagrange multiplier for the capital constraint (21) in the paper also has a payoff similar to a European call option

$$\lambda_1^*(t) = \begin{cases} 2 \tau \left[ \chi M(\tau, X) + M(2\tau, X) \right] & \text{for } r^K_{\tau,t} > r^*_{\tau,t} \\ 0 & \text{for } r^K_{\tau,t} \leq r^*_{\tau,t}. \end{cases}$$

The essential difference from (57) in the paper is that the slope of the payoff is now influenced by the expectation of the marginal investor’s intertemporal rate of substitution, \( \chi M(\tau, X) + M(2\tau, X) \), as well as the weights on two period loans in the capital constraint. This result can be seen by comparing this with (57) in the paper. As a result, the relative payoff between the two and one period loans satisfies

$$\frac{r^K_{2r,t} - r^*_{2r,t}}{r^K_{\tau,t} - r^*_{\tau,t}} = \frac{\chi M(\tau, X) + M(2\tau, X)}{M(\tau, X) + M(2\tau, X)} \chi.$$
Thus, the relative payoffs on two and one period loans is dependent on the stochastic discount factor for the marginal investor, the response of the demand for loans to the two period interest rate, and the weights on two and one period loans in the capital constraint.

4 The Capital Option Value under Capital Constraint

Now that we have the probability distributions (17) we can evaluate the expected marginal value of the loan desk’s capital (65) in the paper. With the solution to this expected marginal value of the loan desk’s capital, we can then determine the optimal amount of the loan desk’s capital using (23) in the paper. To find this option value we use the logic for solving the Black-Scholes option pricing formula.

To illustrate these calculations we start with the counter cyclical buffer. Using (11) in the paper the counter cyclical buffer is given by

$$\frac{P_{\tau,\tau}}{P_{\bar{\tau},\tau}} = \exp\left\{ b_\tau \left[ e^{-A^\tau(\tau-t)}(X - \bar{X}) + Y \right]\right\} = \mathcal{P}(\tau, X) \exp \left\{ b_\tau Y \right\}$$

(25)

The last step uses the fact that $Y$ has a normal distribution with mean 0 and variance $b_\tau K(\tau)b_\tau'$ so that $\frac{P_{\tau,\tau}}{P_{\bar{\tau},\tau}}$ has a standard normal distribution after adjusting for the variance.

As a result, the counter cyclical buffer is positive when

$$Y < e^{-A^\tau(\tau-t)}(X - \bar{X}),$$

since $b_\tau < 0$.

The counter cyclical buffer applies whether or not the liquidity or the capital constraint binds. As a result, we can calculate the cost of the counter cyclical buffer. We need the probability distribution for $\frac{M_{\tau,\tau}}{M_{\bar{\tau},\bar{\tau}}}$. Following the derivation of the forward Kolmogorov equation the two terms have a normal distribution in (17).

$$\exp \left\{ -\frac{1}{2} Y' A_3(\tau)^{-1} Y + b_\tau Y \right\} = \exp \left\{ -\frac{1}{2} (Y - A_3(\tau)b_\tau')' A_3(\tau)^{-1} (Y - A_3(\tau)b_\tau') \right\} + \frac{1}{2} b_\tau A_3(\tau)b_\tau'$$

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As a result we have

$$\frac{M_{2\tau,t}}{M_{t,t}} \left( \frac{\bar{P}_{\tau,s}}{P_{\tau,s}} - 1 \right)^+ = \left( \left( \mathcal{P}(\tau, X) \mathcal{M}(2\tau, X) p_M(t, X, 2\tau, Y) \exp \left\{ b_r Y \right\} \right. \right.$$

$$\left. \left. - \mathcal{M}(2\tau, X) p_M(t, X, 2\tau, Y) \right) \right)^+$$

$$= \frac{1}{\sqrt{(2\pi)^N \det(A_3(\tau))}} \left( \left( \mathcal{P}(\tau, X) \mathcal{M}(2\tau, X) \exp \left\{ \frac{1}{2} b_r A_3(\tau)^{-1} b'_r \right\} \right. \right.$$

$$\times \exp \left\{ - \frac{1}{2} (Y - A_3(\tau)b'_r)' A_3(\tau)^{-1} (Y - A_3(\tau)b'_r) \right\} \right. \left. \left. - \mathcal{M}(2\tau, X) \exp \left\{ - \frac{1}{2} Y' A_3(\tau)^{-1} Y \right\} \right) \right)^+.$$

Let $\rho_b$ for given $X$ be defined by

$$\mathcal{P}(\tau, X) \mathcal{M}(2\tau, X) \exp \left\{ \frac{1}{2} b_r A_3(\tau)b'_r \right\} \exp \left\{ - \frac{1}{2} \left( \rho_b - A_3(\tau)b'_r \right)' A_3(\tau)^{-1} \left( \rho_b - A_3(\tau)b'_r \right) \right\}$$

$$- \mathcal{M}(2\tau, X) \exp \left\{ - \frac{1}{2} \rho'_b A_3(\tau)^{-1} \rho_b \right\} = 0 \Rightarrow \exp \left\{ b_r e^{-A_{\tau-t}(X - \bar{X}) + b_r \rho_b} \right\} - 1 = 0$$

$$\Rightarrow \rho_b = e^{-A_{\tau-t}(X - \bar{X})}.$$

Let $A_3(\tau) = \Sigma_M \Sigma'_M$. Thus, the option value of the counter cyclical buffer is

$$c_b E_t \left( \frac{M_{2\tau,t}}{M_{t,t}} \left( \frac{\bar{P}_{\tau,s}}{P_{\tau,s}} - 1 \right)^+ \right)$$

$$= c_b \mathcal{M}(2\tau, X) \left( \mathcal{P}(\tau, X) \exp \left\{ \frac{1}{2} b'_r A_3(\tau)^{-1} b_r \right\} (1 - \Phi \left( \Sigma_{1M}^{-1} (\rho_b - (\Sigma_M \Sigma'_M) b_r) \right)) \right.$$}

$$\left. - (1 - \Phi \left( \Sigma_{1M}^{-1} \rho_b) \right) \right).$$

This corresponds to (25) in the paper.

Here, define the probability

$$Pr \{ Z > \rho \} \equiv \frac{1}{\sqrt{(2\pi)^N}} \int_{\rho}^{\infty} e^{-\frac{1}{2} Z'Z} dZ \equiv \Phi(\rho),$$

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so that
\[
\frac{\partial \Phi (\rho)}{\partial \rho} = - \frac{1}{\sqrt{(2\pi)^N}} e^{-\frac{1}{2} \rho^2}. 
\]

In the rest of the derivation we take the counter cyclical buffer as given by CCB, since its marginal cost is known given current information.

### 4.1 Option value of capital constraint

The option value of capital for the loan desk is dependent on the marginal value of the capital of the loan desk in the future. This marginal value of capital is

\[
\frac{\partial V}{\partial K_{jL}(t)} = \mathcal{M}(\tau, X) \left[ r^D(t) + \lambda_1^*(t) + \lambda_2^*(t) \right] \tau + E_t \left[ \frac{M_{2\tau,t}}{M_{t,t}} \frac{\partial V}{\partial K_{jL}(t)} \right] 
\]

\[
= \mathcal{M}(\tau, X) \left[ r^D(t) + \lambda_1^*(t) + \lambda_2^*(t) \right] \tau + \mathcal{M}(2\tau, X) E_t \left[ p_M(2\tau,Y) \frac{\partial V}{\partial K_{L}^j(t)} \right] 
\]

\[
= \mathcal{M}(\tau, X) \left[ r^D(t) + \lambda_1^*(t) + \lambda_2^*(t) + (\chi - 1) \right] \tau 
\]

The first step uses the property of the stochastic discount factor which divides it into an expected and random component. The second step uses (44) from the paper. As a result, the expected marginal value of capital for the loan desk is

\[
\mathcal{M}(2\tau,X)p_M(2\tau,Y) \frac{\partial V}{\partial K_{L}^j(t + 1)} = \mathcal{M}(2\tau,X)p_M(2\tau,Y) 
\]

\[
\times \left[ r^D(t + 1) + \lambda_1^*(t + 1) + \lambda_2^*(t + 1) + (\chi - 1) \right] \tau, 
\]

where we have used the properties of the stochastic discount factor from \( t \) to \( t + 2\tau \).

We start with the expression for the expected marginal value of the loan desk’s capital under the capital constraint in (65) of the paper. This term is dependent on the term

\[
E_t \left\{ p_M(2\tau,Y) \left[ \frac{1}{\kappa_L} \left( r_{j\tau,t+\tau}^{j*} - r_{j\tau,t+\tau}^{j} \right) \right] \right\}. 
\]

We will bring back the constant \( \frac{2\chi}{\kappa_L} \mathcal{M}(2\tau,X) \) after the derivation.

Recall from (55) in the paper

\[
r_{j\tau,t}^{j*} = \frac{1}{2} \left( c^j + r_{\tau,t}^{D} \right) + \frac{\gamma_0^j + \sigma_0^j \varepsilon_{\tau,t}^j}{2 \left( \gamma_1^j - \sigma_1^j \varepsilon_{\tau,t}^j \right)}, \quad (33)
\]
As a result, the Lagrange multiplier (57) in the paper is
\[
\lambda^j_{\tau,t} = 2 \frac{\lambda}{\kappa_L} \left[ \frac{1}{(\gamma^j_{0,\tau} - \sigma_0)} \left[ (\gamma^j_{0,\tau} + \sigma_0) - L^j_{\kappa,L} \right] - \frac{1}{2} \left( c^j + r^D_{\tau,t} \right) - \frac{\gamma^j_{0,\tau} + \sigma_0}{2} \right] \]
and the loans subject to the capital constraint (21) in the paper yields
\[
L^j_{\kappa,L} \kappa_0 = 1 \left[ K^j(t + \tau) - \kappa_T \xi K^j_M(t + \tau) - \kappa_L \left( L^j_{\kappa,L} + L^j_{\kappa,R} \right) \right.
- \left. \sigma_0 \left( P^*_{\tau,t} - 1 \right)^+ \right].
(35)

As a result, the Lagrange multiplier (57) in the paper is
\[
\lambda^j_{\tau,t} = 2 \frac{\lambda}{\kappa_L} \left[ \frac{1}{(\gamma^j_{0,\tau} - \sigma_0)} \left[ (\gamma^j_{0,\tau} + \sigma_0) - L^j_{\kappa,L} \right] - \frac{1}{2} \left( c^j + r^D_{\tau,t} \right) - \frac{\gamma^j_{0,\tau} + \sigma_0}{2} \right] \]
and the loans subject to the capital constraint (21) in the paper yields
\[
L^j_{\kappa,L} \kappa_0 = 1 \left[ K^j(t + \tau) - \kappa_T \xi K^j_M(t + \tau) - \kappa_L \left( L^j_{\kappa,L} + L^j_{\kappa,R} \right) \right.
- \left. \sigma_0 \left( P^*_{\tau,t} - 1 \right)^+ \right].
(35)

As a result, the Lagrange multiplier (57) in the paper is
\[
\lambda^j_{\tau,t} = 2 \frac{\lambda}{\kappa_L} \left[ \frac{1}{(\gamma^j_{0,\tau} - \sigma_0)} \left[ (\gamma^j_{0,\tau} + \sigma_0) - L^j_{\kappa,L} \right] - \frac{1}{2} \left( c^j + r^D_{\tau,t} \right) - \frac{\gamma^j_{0,\tau} + \sigma_0}{2} \right] \]
and the loans subject to the capital constraint (21) in the paper yields
\[
L^j_{\kappa,L} \kappa_0 = 1 \left[ K^j(t + \tau) - \kappa_T \xi K^j_M(t + \tau) - \kappa_L \left( L^j_{\kappa,L} + L^j_{\kappa,R} \right) \right.
- \left. \sigma_0 \left( P^*_{\tau,t} - 1 \right)^+ \right].
(35)

The option payoff for the capital constraint using (55) and (56) in the paper is
\[
\frac{2\lambda}{\kappa_L} p_M(2\tau,Y) \left( r^i_{\tau,t+\tau} - r^i_{\tau,t+\tau} \right)^+
= \frac{2\lambda}{\kappa_L} p_M(2\tau,Y) \left( \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \right) - L^j_{\kappa,L} \right. - \left. \frac{1}{2} \left( \gamma^j_{0,\tau} - \sigma_0 \right) \right) \left( c^j + r^D_{\tau,t+\tau} \right)^+.
(36)

This imposes a bound on the constant which determines whether or not the constraint binds.
\[
\frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \right) > L^j_{\kappa,L} + \frac{1}{2} \left( \gamma^j_{0,\tau} - \sigma_0 \right) \left( c^j + r^D_{\tau,t+\tau} \right).
(37)

The conditional expectation of the counter cyclical buffer in the capital constraint (23) in
the paper was found in (29). This is the simplest term in (36) so that it illustrates how to calculate
the whole expression (36).
We continue with the calculation of the probability distribution for the option payoff (36). We now use the rules for future capital of the trading desk (36) in the paper and the factors (4) in the paper. We also use the linear rules for the deposit rate (51) and bank reserves (52) in the paper.

\[
p_M(2\tau, Y) \left[ \frac{1}{2} \left( \gamma_{0, \tau} + \sigma_0 \varepsilon_{\tau,t}^j \right) - L_{\kappa,t}^j - \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) (c^j + r_{\tau,t}^D) \right]
= \frac{CCB}{M(2\tau, X)} + p_M(2\tau, Y) \left[ -\frac{1}{\kappa_L} K_L^j(t + \tau) - \frac{1}{\kappa_L} (1 - \kappa_T \xi) K_M^j(t + \tau) + (L_{2\tau,t+\tau}^j + L_{2\tau,t}^j) \right.
+ \frac{1}{2} \left( \gamma_{0, \tau} + \sigma_0 \varepsilon_{\tau,t}^j \right) - \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) c^j - \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) (d_0 + d_1 X(t + \tau)) \left.ight]
= \frac{CCB}{M(2\tau, X)} + \left\{ \frac{1}{2} \left( \gamma_{0, \tau} + \sigma_0 \varepsilon_{\tau,t}^j \right) - \frac{1}{\kappa_L} K_L^j(t + \tau) + (L_{2\tau,t+\tau}^j + L_{2\tau,t}^j) \right.
- \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) d_1 \left[ e^{-A^p \tau} X(t) + \left( I - e^{-A^p \tau} \right) (A^p)^{-1} \gamma^p \right] \left. \right\} p_M(2\tau, Y)
- \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) d_1 p_M(2\tau, Y) Y
- \frac{1}{\kappa_L} (1 - \kappa_T \xi) K_M^j(t) p_M(2\tau, Y) \exp \left\{ \int_t^{t+\tau} \left[ C_1(\tau) + C_2(\tau) X(s) + \frac{1}{2} X(s)^2 C_3(\tau) X(s) \right] ds \right.
\left. + \int_t^{t+\tau} \left[ C_4(\tau) + C_5(\tau) X(s) \right] ds \right\}
= \frac{CCB}{M(2\tau, X)} + \left\{ \frac{1}{2} \left( \gamma_{0, \tau} + \sigma_0 \varepsilon_{\tau,t}^j \right) - \frac{1}{\kappa_L} K_L^j(t + \tau) + (L_{2\tau,t+\tau}^j + L_{2\tau,t}^j) \right.
- \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) \left( c^j + d_0 + d_1 \mu(\tau, X) \right) \left. \right\} \exp \left\{ -\frac{1}{2} \frac{Y^t \sigma_M(2\tau)^{-1} Y}{(2\pi)^N \det(\sigma_M(2\tau))} \right\}
- \frac{1}{2} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) d_1 \frac{\exp \left\{ -\frac{1}{2} Y^t (\sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1}) Y \right\}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}}
- \frac{1}{\kappa_L} \left( \gamma_{1, \tau} - \sigma_1 \varepsilon_{\tau,t}^j \right) \frac{\exp \left\{ -\frac{1}{2} Y^t (\sigma_M(2\tau)^{-1} + \sigma_K(\tau)^{-1}) Y \right\}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}}.
\]

In the first equality we substitute \( CCB \) for the counter cyclical buffer, the critical loans (35), the deposit rate (51) and bank reserves (52) in the paper. In the second step we
substitute in the rules for future capital of the trading desk (36) in the paper and the factors (4) in the paper. In the final step we use the appropriate probability rules (17).

We also used the definition

\[ \mu(\tau, X) \equiv e^{-A^p\tau}X + \left( I - e^{-A^p\tau} \right) (A^p)^{-1} \gamma^p \]

We want to evaluate the option value from equation (65) in the paper.

\[ \frac{2X}{\kappa L} E_t \left[ p_M(2\tau, Y) \left( r_{j^2}^{j^1}_{\tau,t+\tau} - r_{j^2}^{j^1}_{\tau,t+\tau} \right) \right] = \frac{2X}{\kappa L} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) \mathcal{M}(2\tau, X) + \frac{2X}{\kappa L} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) \int_{-\infty}^{\infty} \left\{ \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{0,\tau} \right) - \frac{1}{\kappa L} K_L^j(t + \tau) + \left( L_{2\tau,t+\tau}^j + L_{2\tau,t}^j \right) \right. \]

\[ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) (c^j + d_0 + d_1 \mu(\tau, X)) \exp \left\{ - \frac{1}{2} Y' \sigma_M(2\tau)^{-1} Y \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} \]

\[ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) d_1 \exp \left\{ - \frac{1}{2} Y' \left( \sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1} \right) Y \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \]}

\[ \left( (1 - \kappa \tau \xi) K_M^j(t) \mathcal{K}(\tau, X) \exp \left\{ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) (c^j + d_0 + d_1 \mu(\tau, X)) \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} \right. \]

Find \( \rho_\kappa \) for each \( \{ K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j \} \) such that

\[ F \left( \rho_\kappa, K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j \right) = \left\{ \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{0,\tau} \right) - \frac{1}{\kappa L} K_L^j(t + \tau) + \left( L_{2\tau,t+\tau}^j + L_{2\tau,t}^j \right) \right\} \exp \left\{ - \frac{1}{2} \rho_\kappa \sigma_M(2\tau)^{-1} \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} \]

\[ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) \left( c^j + d_0 + d_1 \mu(\tau, X) \right) \exp \left\{ - \frac{1}{2} \rho_\kappa \sigma_M(2\tau)^{-1} \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \]

\[ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) d_1 \exp \left\{ - \frac{1}{2} \rho_\kappa \sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1} \right\} \rho_\kappa \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \]}

\[ - (1 - \kappa \tau \xi) K_M^j(t) \mathcal{K}(\tau, X) \exp \left\{ - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{1,\tau} \right) (c^j + d_0 + d_1 \mu(\tau, X)) \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}} \] = 0.
Cancel the common terms to find

\[
F (\rho, K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j) = \left\{ \frac{1}{2} (\gamma_{0,\tau} + \sigma_0 \varepsilon^j_{\tau,t}) - \frac{1}{\kappa_L} K_L^j(t + \tau) + (L_{2\tau,t+\tau}^j + L_{2\tau,t}^j) \right. \\
- \frac{1}{2} (\gamma_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t}) (c^j + d_0 + d_1 \mu(\tau, X)) \left\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} \\
- \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t} \right) d_1 \exp \left\{ -\frac{1}{2} \rho_\kappa' \sigma Y(\tau)^{-1} \rho_\kappa \right\} \\
\frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} = 0. \tag{39}
\]

This result leads to the equation (67) in the paper.

This relation is nonlinear in \(\rho_k\), because of the terms like

\[
f(\rho_\kappa, K_M) = \exp \left\{ -\frac{1}{2} \rho_\kappa' \sigma K(\tau)^{-1} \rho_\kappa \right\}. \tag{40}
\]

However, we do know

\[
\frac{\partial F (\rho_\kappa, K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j)}{\partial \rho_\kappa} = \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t} \right) d_1 \exp \left\{ -\frac{1}{2} \rho_\kappa' \sigma Y(\tau)^{-1} \rho_\kappa \right\} \frac{2 \rho_\kappa' \sigma Y(\tau)^{-1}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \\
+ \frac{(1 - \kappa_T \xi)}{\kappa_L} K_M^j(t) K(\tau, X) \exp \left\{ -\frac{1}{2} \rho_\kappa' \sigma K(\tau)^{-1} \rho_\kappa \right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}} \geq 0, \tag{41}
\]

so the function is always increasing in the cutoff. In addition, it reaches its minimum at \(\rho_\kappa = 0\).

If

\[
F (0, K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j) < 0, \tag{42}
\]

then there exists exactly two solutions \(\rho_\kappa\). Otherwise the capital constraint is always binding, so that you just take the expected value without a cutoff.

We have the partial derivative for loan desk manager’s future capital

\[
\frac{\partial F (\rho_\kappa, K_L^j(t + \tau), K_M^j(t), X, \varepsilon_i^j)}{\partial K_L^j} = -\frac{1}{\kappa_L} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} < 0. \tag{43}
\]
We have the partial derivative for the trading desk’s capital

$$\frac{\partial F(\rho, K^j_L(t + \tau), K^j_M(t), X, \varepsilon^j_i)}{\partial K^j_M(t)} = -\frac{(1 - \kappa_T \xi)}{\kappa L} \frac{\exp\left\{-\frac{1}{2} \rho' \sigma K(\tau)^{-1} \rho\right\}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}} < 0. \quad (44)$$

We have the partial derivative for the yield curve factors.

$$\frac{\partial F(\rho, K^j_L(t + \tau), K^j_M(t), X, \varepsilon^j_i)}{\partial X} = -\frac{1}{2} \left(\gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t}\right) d_1 e^{-A^\tau} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}}$$

$$+ \frac{(1 - \kappa_T \xi)}{\kappa L} K^j_M(t) K(\tau, X) (\sigma_K(\tau))^{-1} \left(X - \mu_K(\tau)\right) \frac{\exp\left\{-\frac{1}{2} \rho' \sigma_K(\tau)^{-1} \rho\right\}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}}. \quad (45)$$

This partial derivative is used to find conditions 1 and 2 on page 37 of the paper.

We have the partial derivative for the loan shocks.

$$\frac{\partial F(\rho, K^j_L(t + \tau), K^j_M(t), X, \varepsilon^j_i)}{\partial \varepsilon^j_i} = \left\{\frac{1}{2} \sigma_0 + \sigma_1 \left(c^j + d_0 + d_1 \mu(\tau, X)\right)\right\} \frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}}$$

$$+ \frac{1}{2} \sigma_1 d_1 \frac{\exp\left\{-\frac{1}{2} \rho' \sigma_Y(\tau)^{-1} \rho\right\}}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} > 0. \quad (46)$$

The total differential is

$$dF(\rho, K^j_L(t + \tau), K^j_M(t), X, \varepsilon^j_i) = F_{\rho \rho} d\rho + F_{K^j_L(t + \tau)} dK^j_L(t + \tau)$$

$$+ F_{K^j_M(t)} dK^j_M(t) + F_X dX + F_{\varepsilon^j_i} d\varepsilon^j_i. \quad (47)$$

It is straight forward to calculate all these derivatives using the Chebfun add on for Matlab.

As a result, we have

$$\frac{\partial \rho}{\partial K^j_L(t + \tau)} = -\frac{F_{K^j_L(t + \tau)}}{F_{\rho \rho}} > 0, \quad \frac{\partial \rho}{\partial K^j_M(t)} = -\frac{F_{K^j_M(t)}}{F_{\rho \rho}} > 0,$$

$$\frac{\partial \rho}{\partial X} = -\frac{F_X}{F_{\rho \rho}}, \quad \text{and} \quad \frac{\partial \rho}{\partial \varepsilon^j_i} = -\frac{F_{\varepsilon^j_i}}{F_{\rho \rho}} < 0. \quad (48)$$

These results are the bases for Proposition 5.3 in the text.
In choosing the optimal capital for the loan desk we need to know how this critical value is affected by changes in the loan desk’s capital stock.

\[
\frac{\partial \rho_\kappa(K_j^i(t + \tau), \Omega_{t,\tau})}{\partial K_j^i(t + \tau)} > 0, \quad \text{and} \quad \frac{\partial^2 \rho_\kappa(K_j^i(t + \tau), \Omega_{t,\tau})}{\partial^2 K_j^i(t + \tau)} < 0.
\]

Thus, the critical function \( \rho_\kappa(K_j^i(t + \tau), \Omega_{t,\tau}) \) must be concave in \( K_j^i(t + \tau) \).

We also want to find the Cholesky decompositions for \( \sigma_M(2\tau) = \Sigma_M \Sigma'_M, \sigma_M(2\tau) + \sigma_Y(\tau) = \Sigma_M \Sigma'_MY \) and \( \sigma_M(2\tau) + \sigma_K(\tau) = \Sigma_M \Sigma'_{MK} \). In this case the option value using (55) and (56) in the paper is

\[
\frac{2\chi}{\kappa_L} M(2\tau, X) E_t \left[ p_M(2\tau, Y) (\gamma^j_{\tau,t+\tau} - r^j_{\tau,t+\tau} + 1)^+ \right] = \frac{2\chi}{\kappa_L} M(2\tau, X) \left\{ \begin{array}{c} CCB(X) \\ M(2\tau, X) \end{array} \right\}
\]

\[
+ \left[ \frac{1}{2} (\gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{\tau,t}) - \frac{1}{\kappa_L} K_j^i(t + \tau) + (L_j^2(t) + L_j^2(t)) \right. \\
- \frac{1}{2} (\gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t}) (c^j + d_0 + d_1 \mu(\tau, X)) \\
- \frac{1}{2} (\gamma^j_{1,\tau} - \sigma_1 \varepsilon^j_{\tau,t}) d_1 \Phi(\Sigma_Y(\tau)^{-1}\rho_\kappa) \sqrt{det(\sigma_M(2\tau))} \\
\left. \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \right] \Phi(\Sigma_{-1}^M \rho_\kappa) \\
- \frac{(1 - \gamma_{T,\xi})}{\kappa_L} K_j^i(t) K(\tau, X) \Phi(\Sigma_K(\tau)^{-1}\rho_\kappa) \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_K(\tau))}} \right].
\]
As a result, we can calculate the Delta for the option value of the loan desk’s capital.

$$\Delta_\kappa = \frac{\partial M(2\tau, X) E_t \left[ \rho M(2\tau, Y) \left[ \frac{r_j^{\kappa}}{r_{\tau+t}^{\kappa}} - r_{\tau+t}^{\kappa} \right] + \right]}{\partial K'_L(t+\tau)} = \frac{2\chi M(2\tau, X)}{\kappa_L \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right)} \left\{ - \frac{1}{K_L} \right\}$$

(50)

$$- \left[ \frac{1}{2} \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right) d_1 \frac{\sqrt{\text{det}(\sigma_M(2\tau))}}{\sqrt{\text{det}(\sigma_M(2\tau) + \sigma_Y(\tau))}} \frac{\partial \Phi(\Sigma_Y(\tau)^{-1} \rho_\kappa)}{\partial \rho_\kappa} \right]$$

$$+ \frac{2\chi M(2\tau, X)}{\kappa_L \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right)} - \frac{CCB(X)}{M(2\tau, X)} + \frac{1}{2} \left( \gamma_0^j + \sigma^j \varepsilon_{\tau,t}^j \right) - \frac{1}{\kappa_L} K_L^j(t+\tau) + (L_2^j + L_2^j)$$

$$- \frac{1}{2} \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right) \left( c^j + d_0 + d_1 \mu(\tau, X) \right)$$

$$- \frac{1}{2} \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right) d_1 \Phi(\Sigma_Y(\tau)^{-1} \rho_\kappa) \frac{\sqrt{\text{det}(\sigma_M(2\tau))}}{\sqrt{\text{det}(\sigma_M(2\tau) + \sigma_Y(\tau))}}$$

$$- \frac{1}{2} \left( \gamma_1^\kappa - \sigma^j \varepsilon_{\tau,t}^j \right) d_1 \Phi(\Sigma_Y(\tau)^{-1} \rho_\kappa) \frac{\sqrt{\text{det}(\sigma_M(2\tau))}}{\sqrt{\text{det}(\sigma_M(2\tau) + \sigma_Y(\tau))}}$$

$$\frac{\partial \Phi(\Sigma_K(\tau)^{-1} \rho_\kappa)}{\partial \rho_\kappa} = - \frac{\exp \left\{ - \frac{1}{2} \rho_\kappa^2 (\sigma_K(\tau)^{-1} \rho_\kappa) \right\}}{\sqrt{(2\pi)^N \text{det}(\sigma_K(\tau))}} (\sigma_K(\tau)^{-1}) \rho_\kappa < 0.$$  

(51)

There are four competing effects of an increase in the capital for the loan desk’s:

1. a direct negative effect on the capital constraint, since more capital lessens the capital constraint.

2. a negative effect through the cumulative probability that the capital constraint binds using the probability distribution for the stochastic discount factor $\Phi(\Sigma_Y(\tau)^{-1} \rho_\kappa)$, by Proposition 5.3.

3. a positive indirect effect through $\Phi(\Sigma_Y(\tau)^{-1} \rho_\kappa)$, the cumulative probability that the capital constraint binds using the probability distributions for the future factors, $Y$ in (5) from the paper.

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4. a positive indirect effect through $\Phi(\Sigma_K(\tau)^{-1}\rho_h)$, the cumulative probabilities that the capital constraint binds using the probability distribution for the capital for the trading desk (35) in the paper.

To satisfy the second order condition for the loan desk’s to issue equity or pay dividends, we must have $\Delta_h < 0$. In addition, to have an interior solution for the issuing of equity or payment of dividends (43) or (44), we make the following assumption:

**Assumption 1**: The absolute value of the sum of effects 1. and 2. is greater than the sum of effects 3. and 4.

Next we want to know how changes in the current capital of the trading desk influences the expected marginal value of capital for the loan desk (65) in the paper $\Delta_{KM} = \partial EMV(X,K^j_L(t),K^j_L(t+\tau))$.

$$\Delta_{KM} = \frac{\partial M(2\tau,X)E_t}{\partial K^j_M(t)} \left[ p_M(2\tau,Y) \left[ r^{j\kappa}_{\tau,t+\tau} - r^{j\kappa}_{\tau,t+\tau} \right] + \right] - \frac{2\chi M(2\tau,X)}{\kappa_L} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) \frac{\partial EMV(X,K^j_L(t),K^j_L(t+\tau))}{\partial K^j_M(t)}$$

$$\times K(\tau,X) \Phi(\Sigma_K(\tau)^{-1}\rho_h) \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_K(\tau))}} \Phi(\Sigma_M^{-1}\rho_h)$$

$$+ \left[ \frac{2\chi}{\kappa_L} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \varepsilon_{1,\tau, t} \right) - \frac{1}{\kappa_L} K^j_L(t + \tau) + \left( L^j_{2,\tau,t+\tau} + L^j_{2,\tau,t} \right) - \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) \left( c^j + d_0 + d_1 \mu(\tau, X) \right) \right]$$

$$- \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) d_1 \Phi(\Sigma_Y(\tau)^{-1}\rho_h) \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_Y(\tau))}}$$

$$- \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) d_1 \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \left[ \frac{\partial EMV(X,K^j_L(t),K^j_L(t+\tau),\Omega_{\tau,t})}{\partial \rho_h} \frac{\partial EMV(X,K^j_L(t),K^j_L(t+\tau),\Omega_{\tau,t})}{\partial K^j_M(t)} \right]$$

$$- \frac{1}{\kappa_L} \left( \gamma^j_{1,\tau} - \sigma_1 \varepsilon_{1,\tau, t} \right) d_1 \frac{\sqrt{det(\sigma_M(2\tau))}}{\sqrt{det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \left[ \frac{\partial EMV(X,K^j_L(t),K^j_L(t+\tau),\Omega_{\tau,t})}{\partial \rho_h} \frac{\partial EMV(X,K^j_L(t),K^j_L(t+\tau),\Omega_{\tau,t})}{\partial K^j_M(t)} \right] \Phi(\Sigma_M^{-1}\rho_h)$$

The four effects of this change are the same as for a change in the capital for the loan desk’s. The key change is that the changes in the three cumulative probabilities for the
stochastic discount factor, yield curve factor and capital for the trading desk now have different qualitative effects, based on the size of \( \frac{\partial \rho_n}{\partial K^j(t+\tau)} > 0 \) relative to \( \frac{\partial \rho_n}{\partial K^j(t)} > 0 \) from Proposition (5.3) in the paper.

The impact of a change in the interest rate factors \( \Delta X = \frac{\partial EMV(X,K^j(t),K^j(t+\tau))}{\partial X} \) is given by

\[
\Delta X = \frac{2\chi}{\kappa_L (\gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j)} \frac{\partial CCB(X)}{\partial X}
\]

\[
+ \frac{2\chi \mathcal{M}(2\tau, X)}{\kappa_L (\gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j)} \left\{ \frac{1}{2} \left( \gamma_{0,\tau} + \sigma_0 \xi_{\tau,t}^j \right) - \frac{1}{\kappa_L} K^j_L(t+\tau) + (L_{2\tau,t+\tau} + L_{2\tau,t}) \right. \\
- \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j \right) (c^j + d_0 + d_1 \mu(\tau, X)) \left. \right\}
\]

\[
\times \left[ \frac{\partial \Phi \left( \Sigma^{-1}_M \rho_n \right)}{\partial \rho_n} \frac{\partial \rho_n (K^j_L(t+\tau), \Omega_{t,\tau})}{\partial X} + (X - \mu_M) \Phi \left( \Sigma^{-1}_M \rho_n \right) \right]
\]

\[
+ \left\{ \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j \right) \frac{\sqrt{\det(\sigma_M(2\tau))}}{\sqrt{\det(\sigma_M(2\tau) + \sigma_K(\tau))}} \frac{\partial \Phi \left( \Sigma^{-1}_Y \rho_n \right)}{\partial \rho_n} \frac{\partial \rho_n (K^j_L(t+\tau), \Omega_{t,\tau})}{\partial X} \right. \\
- \frac{(1 - \kappa_T \xi)}{\kappa_L} K^j_M(t) \mathcal{K}(\tau, X) \Phi \left( \Sigma_K(\tau)^{-1} \rho_n \right) \frac{\sqrt{\det(\sigma_M(2\tau))}}{\sqrt{\det(\sigma_M(2\tau) + \sigma_K(\tau))}} \frac{\partial \Phi \left( \Sigma^{-1}_K \rho_n \right)}{\partial \rho_n} \frac{\partial \rho_n (K^j_L(t+\tau), \Omega_{t,\tau})}{\partial X} \right. \\
- \frac{(1 - \kappa_T \xi)}{\kappa_L} K^j_M(t) \mathcal{K}(\tau, X) (X - \mu_K) \Phi \left( \Sigma_K(\tau)^{-1} \rho_n \right) \\
- \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j \right) \frac{\sqrt{\det(\sigma_M(2\tau))}}{\sqrt{\det(\sigma_M(2\tau) + \sigma_K(\tau))}} \frac{\partial \Phi \left( \Sigma^{-1}_M \rho_n \right)}{\partial \rho_n} \frac{\partial \rho_n (K^j_L(t+\tau), \Omega_{t,\tau})}{\partial X} \right. \\
- \frac{1}{2} \left( \gamma_{1,\tau} - \sigma_1 \xi_{\tau,t}^j \right) d_1 e^{-\lambda P(\tau-t)} \left. \right\} \Phi \left( \Sigma^{-1}_M \rho_n \right)
\]

(53)

There are four effects of changes in the yield curve factors.

1. a direct negative effect through the decrease in the cost of the counter cyclical buffer by (29).

2. a negative effect through a change in the conditional expectation of the gross rate of capital growth for the trading desk, \( \mathcal{K}(\tau, X) \), for \( X > \mu_K \), and the change in the future deposit rate from (51) in the paper resulting from the change in the expected value of the yield curve factors by (4) in the paper.

3. a positive effect through a change in the cumulative probability distribution, \( \Phi \left( \Sigma^{-1}_M \rho_n \right) \)
for the yield curve factors, and the gross growth rate of trading desk’s capital, \( \Phi \left( \Sigma_K^{-1} \rho_n \right) \), by Proposition 5.3 in the paper.

4. a positive effect through the cumulative probability distribution, \( \Phi \left( \Sigma_M^{-1} \rho_n \right) \), and the conditional expected value, \( M(2\tau, X) \), of the stochastic discount factor, when \( X > \mu_M > \mu_K \) by Proposition 5.3 in the paper.

### 4.2 The Capital Option Value under Liquidity Constraint

We also need to find the option value of the liquidity constraint to complete the calculations in (65) in the paper. Recall from (22) from the paper.

\[
L^j_{\tau,t+\tau} = \frac{1}{\alpha_\tau} \left[ K^j_L(t+\tau) + (1 - \alpha T \xi) K^j_M(t+\tau) + \alpha_R R^j_{\tau,t+\tau} - \alpha_2 \left( L^j_{\tau,t+\tau} + L^j_{\tau,t} \right) \right]
\]

The following calculations are similar to those for the capital constraint. We use the rules for future capital of the trading desk (36) in the paper and the yield curve factors (4) in the paper.

The payoff on the option is

\[
p_M(2\tau, Y) \left[ L^j_{\tau,t+\tau} - \frac{1}{2} (\gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t}) - \frac{1}{2} (-\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}) \right] = p_M(2\tau, Y) \left[ \frac{1}{\alpha_\tau} K^j_L(t + \tau) \right.
\]

\[
+ \frac{1}{\alpha_\tau} (1 - \alpha T \xi) K^j_M(t + \tau) + \frac{1}{\alpha_\tau} \alpha_R (r_0 + r_1 X(t + \tau)) - \frac{1}{\alpha_\tau} \alpha_2 \left( L^j_{\tau,t+\tau} + L^j_{\tau,t} \right)
\]

\[
- \frac{1}{2} (\gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t}) - \frac{1}{2} (-\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}) c^j - \frac{1}{2} \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t} \right) \left( d_0 + d_1 X(t + \tau) \right) \]

\[
= \left\{ \frac{1}{\alpha_\tau} K^j_L(t + \tau) - \frac{1}{\alpha_\tau} \alpha_2 \left( L^j_{\tau,t+\tau} + L^j_{\tau,t} \right) + \frac{1}{\alpha_\tau} \alpha_R r_0 - \frac{1}{2} \left( \gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t} \right) - \frac{1}{2} \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t} \right) c^j + d_0 \right\} p_M(2\tau, Y)
\]

\[
+ \left[ \frac{1}{\alpha_\tau} \alpha_R r_1 - \frac{1}{2} \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t} \right) d_1 \right] p_M(2\tau, Y) Y_{\tau}
\]

\[
+ \frac{1}{\alpha_\tau} (1 - \alpha T \xi) K^j_M(t) p_M(2\tau, Y) \exp \left\{ \int_{t}^{t+\tau} \left[ C_1(\tau) + C_2(\tau) X(s) + \frac{1}{2} X(s)' C_3(\tau) X(s) \right] ds \right\}
\]

\[
+ \int_{t}^{t+\tau} \left[ C_4(\tau) + C_5(\tau) X(s) \right] d\epsilon_s \}
\]
\[
\begin{align*}
&= \left\{ \frac{1}{\alpha_\tau} K^j_L(t + \tau) - \frac{\alpha_2}{\alpha_\tau} \left( L^j_{2,\tau,t+\tau} + L^j_{2,t+\tau} \right) - \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{\tau,t} \right) - \frac{1}{2} \left( -\gamma^j_{1,\tau} + \sigma_1 \varepsilon^j_{\tau,t} \right) \left( c^j + d_0 + d_1 \mu(X) \right) \\
&+ \frac{\alpha_R}{\alpha_\tau} \left( r_0 + r_1 \mu(X) \right) \right\} \exp \left\{ -\frac{1}{2} Y' \sigma_M(2\tau)^{-1} Y \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau))} \\
&+ \left[ \frac{\alpha_R}{\alpha_\tau} r_1 - \frac{1}{2} \left( -\gamma^j_{1,\tau} + \sigma_1 \varepsilon^j_{\tau,t} \right) d_1 \right] \exp \left\{ -\frac{1}{2} Y' \left( \sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1} Y \right) \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))} \\
&+ \left( 1 - \alpha_T \xi \right) K^j_M(t) \mathcal{K}(\tau, X) \exp \left\{ -\frac{1}{2} Y' \left( \sigma_M(2\tau)^{-1} + \sigma_K(\tau)^{-1} Y \right) \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))}.
\end{align*}
\]

The option value of the liquidity constraint becomes

\[
E_t \left[ p_M(2\tau, Y) \left[ r_{i,t+\tau}^{j} - r_{i,t+\tau}^{j*} \right]^+ \right] = \frac{2\chi}{\alpha_\tau} \left( -\gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{\tau,t} \right) \int^\infty_{-\infty} \left\{ \frac{1}{\alpha_\tau} K^j_L(t + \tau) - \frac{\alpha_2}{\alpha_\tau} \left( L^j_{2,\tau,t+\tau} + L^j_{2,t+\tau} \right) \right. \\
\left. - \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0 \varepsilon^j_{\tau,t} \right) - \frac{1}{2} \left( -\gamma^j_{1,\tau} + \sigma_1 \varepsilon^j_{\tau,t} \right) \left( c^j + d_0 + d_1 \mu(X) \right) + \frac{\alpha_R}{\alpha_\tau} \left( r_0 + r_1 \mu(X) \right) \right\}
\times \exp \left\{ -\frac{1}{2} Y' \sigma_M(2\tau)^{-1} Y \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau))} \\
+ \left[ \frac{\alpha_R}{\alpha_\tau} r_1 - \frac{1}{2} \left( -\gamma^j_{1,\tau} + \sigma_1 \varepsilon^j_{\tau,t} \right) d_1 \right] \exp \left\{ -\frac{1}{2} Y' \left( \sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1} Y \right) \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))} \\
+ \left( 1 - \alpha_T \xi \right) K^j_M(t) \mathcal{K}(\tau, X) \exp \left\{ -\frac{1}{2} Y' \left( \sigma_M(2\tau)^{-1} + \sigma_K(\tau)^{-1} Y \right) \right\} \sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_K(\tau))} \right\} dY.
\]
Find \( \rho \) for each \( \{K_L^j(t + \tau), K_M^j(t), X, \varepsilon^j_i\} \) such that

\[
G (\rho, K_L^j(t + \tau), K_M^j(t + \tau), K_M^j(t), X, \varepsilon^j_i) = \left\{ \frac{1}{\alpha} K_L^j(t + \tau) - \frac{\alpha_2}{\alpha_1} (L_{2r,t}^j + L_{2r,t}^j) \right. \\
- \frac{1}{2} \left( \gamma_{0,\tau} + \sigma_0 \varepsilon_{r,t}^j \right) - \frac{1}{2} \left( - \gamma_{1,\tau} + \sigma_1 \varepsilon_{r,t}^j \right) \left( c^j + d_0 + d_1 \mu(X) \right) + \frac{\alpha}{\alpha_1} (r_0 + r_1 \mu(X)) \right\}
\]

(54)

\[
\frac{1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau))}} \exp \left\{ - \frac{1}{2} \rho^i [\sigma_M(2\tau)^{-1} \rho^i] \right\}
\]

\[
+ \frac{\alpha_2 r_1 - \frac{1}{2} \left( - \gamma_{1,\tau} + \sigma_1 \varepsilon_{r,t}^j \right) d_1}{\sqrt{(2\pi)^N \det(\sigma_M(2\tau) + \sigma_Y(\tau))}} \exp \left\{ - \frac{1}{2} \rho^i [\sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1}] \rho^i \right\}
\]

(55)

The Greeks \( \Delta_l \) and \( \Gamma_l \) are the same as in the case of the capital constraint with \( \kappa \) replaced by \( l \).

As a result the option value of capital is given by

\[
\mathcal{M}(2\tau, X) E_t \left[ p_{M}(2\tau, Y) \left[ r_{\tau,t}^j - r_{\tau,t}^* \right]^j \right] = \frac{2\chi \mathcal{M}(2\tau, X)}{\alpha} \left\{ \left( \frac{1}{2} \left( \gamma_{0,\tau} + \sigma_0 \varepsilon_{r,t}^j \right) \right) \right. \\
- \frac{1}{\alpha} K_L^j(t + \tau) + \frac{\alpha_2}{\alpha_1} (L_{2r,t}^j + L_{2r,t}^j) - \frac{1}{2} \left( \gamma_{1,\tau} + \sigma_1 \varepsilon_{r,t}^j \right) \left( c^j + d_0 + d_1 \mu(X) \right) \\
- \frac{\alpha}{\alpha_1} (r_0 + r_1 \mu(X)) \right\} \times \phi \left( \Sigma_{M}^{-1} \rho^i \right) \left[ \frac{\alpha_2 r_1 - \frac{1}{2} \left( - \gamma_{1,\tau} + \sigma_1 \varepsilon_{r,t}^j \right) d_1}{\sigma_M(2\tau)^{-1} + \sigma_Y(\tau)^{-1}} \right] \phi \left( \Sigma_{M}(\tau)^{-1} + \Sigma_{K}(\tau)^{-1} \right) \rho^i \\
+ \frac{1}{\alpha_1} \left( \frac{1}{2} \left( \gamma_{0,\tau} + \sigma_0 \varepsilon_{r,t}^j \right) \right) \phi \left( \Sigma_{M}(\tau)^{-1} + \Sigma_{K}(\tau)^{-1} \right) \rho^i 
\]

(56)

The option value of the loan desk’s capital under both the capital (49) and liquidity constraints (56) have the same functional form when one uses the parameters of each constraint. Thus, the comparative statics for each option value are the same as in the previous subsection.

The last term we have to calculate for the expected marginal value of the loan desk’s capital is the marginal effect of the deposit rate.

\[
E_t \left[ p_{M}(2\tau, Y) \left[ r^D(t + \tau) \right] \right] = E_t \left[ p_{M}(2\tau, Y) \left( d_0 + d_1 X(t + \tau) \right) \right] \\
= E_t \left[ p_{M}(2\tau, Y) \left( d_0 + d_1 \left[ e^{-At} X(t) - \left( I - e^{-At} \right) (A)^{-1} \gamma^p + \int_t^\tau e^{-Ap(\tau-v)} \Sigma X d\nu \right] \right) \right] \\
= (d_0 + d_1 \mu(X) + 1) \tau.
\]
In the first step we use the expression for the deposit rate (51) in the paper. In the second step we use the solution for the factors (4) and (5) in the paper. The third step separates the two random expressions. The last step applies the intertemporal rate of substitution (10) from the paper and the price of an \( \tau \) period zero coupon bond (11) from the paper to evaluate the first term in terms of current factors.

We can now put the three terms together. For any change over one period we have

\[
\mathcal{M}(2\tau, X) E_t \left\{ p_M(2\tau, Y) \left( r^D(t + \tau) \tau + 2\chi \max \left\{ \frac{1}{\alpha_r} \left( r^j_{\tau,t+\tau} - r^j_{\tau,t+\tau} \right)^+; \frac{1}{\kappa_L} \left( r^i_{\tau,t+\tau} - r^i_{\tau,t+\tau} \right)^+ \right\} + (\chi - 1) \tau \right\} = (d_0 + d_1\mu(X) + 1) \mathcal{M}(2\tau, X)\tau
\]

\[
+ \sum_{i=1}^{S} (\varepsilon^j_{t,t} = \varepsilon^i_{t}) \frac{2\chi \mathcal{M}(2\tau, X)}{(\gamma^j_{1,\tau} - \sigma_1\varepsilon^j_{t})} \max \left\{ \frac{1}{\alpha_r} \left\{ \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0\varepsilon^j_{t} \right) \right\} - \frac{1}{\alpha_r} K^j_L(t + \tau)
\]

\[
+ \frac{\alpha_r}{\alpha_r} \left( K^j_M(t) \Phi(\Sigma^{-1}_M \rho_t) - \frac{1}{2} \left( \gamma^j_{1,\tau} + \sigma_1\varepsilon^j_{t} \right) \right) - \frac{\alpha_R}{\alpha_r} \left( r_0 + \frac{1}{\alpha_r} \right) \Phi(\Sigma^{-1}_M \rho_t)
\]

\[
+ \frac{1}{\alpha_r} \left( CCB(X) \right) \frac{1}{\kappa_L} \mathcal{M}(2\tau, X) + \frac{1}{\kappa_L} \left\{ \frac{1}{2} \left( \gamma^j_{0,\tau} + \sigma_0\varepsilon^j_{t} \right) \right\} - \frac{1}{\kappa_L} K^j_L(t + \tau) + (L^j_{2\tau,t+\tau} + L^j_{2\tau,t})
\]

\[
- \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1\varepsilon^j_{t} \right) \left( c^j + d_0 + d_1\mu(t, X) \right) \right\} \Phi(\Sigma^{-1}_M \rho_n)
\]

\[
- \frac{1}{2} \left( \gamma^j_{1,\tau} - \sigma_1\varepsilon^j_{t} \right) d_1 \mathcal{M}(2\tau, X) \Phi((\Sigma_M(\tau)^{-1} + \Sigma_K(\tau)^{-1}) \rho_n)
\]

\[
- \frac{1}{\kappa_L} \left( \Sigma^j_M(\tau)^{-1} + \Sigma_K(\tau)^{-1} \right) \rho_n \}
\]

Here \( t \) can be any time period. Also, the shock to loan demand \( \varepsilon^j_{t} \) has a discrete distribution with \( S \) values \( \varepsilon^j_{t} \). This corresponds to (75) in the text when the two period loans are ignored. The \( \Delta \) in the text corresponds to either \( \Delta_\kappa \) or \( \Delta_t \) depending on which constraint (22) or (23) is binding. The same is true for (53) and (50) for the change in the interest rate factors or the trading desk’s capital stock.
5 Choice of Capital for the Trading and Loan Desks

We can now discuss how the COO chooses the amount of future capital for the loan desk. We start in a terminal period \( t + (n - 1)\tau + \tau \) and choose capital for the last period based on the first order condition for choosing new issues of capital or payment of dividends.\(^3\) We use \((\chi - 1)\) since most of the time a bank pays dividends but seldom issues equity.

As a result, the capital allocated to the loan desk for time \( n\tau \) under (44) in the paper is given by\(^4\)

\[
\mathcal{M}(\tau, X)(\chi - 1)\tau - (d_0 + d_1\mu(X(t + (n - 1)\tau)) + 1)\mathcal{M}(2\tau, X(t + (n - 1)\tau))\tau
- \sum_{i=1}^{S} Pr \left( \varepsilon_{n\tau,t}^j = \varepsilon_i^j \right) \frac{2\chi\mathcal{M}(2\tau, X(t + (n - 1)\tau))}{(\gamma_1^j - \sigma_1\varepsilon_i^j)} \left\{ \frac{1}{\kappa_L} \mathcal{M}(2\tau, X(t + (n - 1)\tau)) + \frac{1}{\kappa_L} \left\{ \frac{1}{2} (\gamma_0^j + \sigma_0\varepsilon_{\tau,t}^j) \right\} \right\}
- \frac{1}{\kappa_L} K_{L}^j(t + \tau) + (L_{2\tau,t+\tau}^j + L_{2\tau,t}^j) - \frac{1}{2} (\gamma_1^j - \sigma_1\varepsilon_{\tau,t}^j) (c^j + d_0 + d_1\mu(\tau, X(t + (n - 1)\tau)))
\times \Phi \left( (\Sigma^{-1}_{M} \rho_{\kappa}((n - 1)\tau)) - \frac{1}{2} (\gamma_1^j - \sigma_1\varepsilon_{\tau,t}^j) d_1 \Phi \left( (\Sigma_{M}(\tau)^{-1} + \Sigma_{Y}(\tau)^{-1}) \rho_{\kappa}((n - 1)\tau) \right) \right)
- \frac{(1 - \kappa T\xi)}{\kappa L} K_{M}^j(t) K(\tau, X(t + (n - 1)\tau)) \Phi \left( (\Sigma_{M}(\tau)^{-1} + \Sigma_{K}(\tau)^{-1}) \rho_{\kappa}((n - 1)\tau) \right) \right\} = 0,
\]

which yields a maximum when

\[ \Delta_\kappa < 0. \]

Here, \( \Delta_\kappa \) is given by (50). This is a fixed point problem that yields the optimal choice of capital in the next to last period, \( K_L^j(t + n\tau) \).

**Theorem 5.1.** The bank’s choice of capital for the loan desk is optimal when (57) holds for the time period \([ (n - 1)\tau, n\tau ] \) and \( \Delta_\kappa < 0. \)

\(^3\)We write the period as \( t + (n - 1)\tau + \tau \) rather than \( t + n\tau \), since in general the loan desk’s capital is chosen in the previous period \( t + (n - j)\tau \) and is available in the next period \( t + (n + 1 - j)\tau \).

\(^4\)To save space we only include the option value of the capital constraint. If the liquidity constraints binds then replace \( l \) with \( \kappa \).
By (30) the marginal value of capital for the lending desk in period \( t + (n - 1)\tau \) is
\[
\frac{\partial V}{\partial K^j_L(t + (n - 1)\tau)} = \mathcal{M}(\tau, X) \left[ r^D(t + (n - 1)\tau)\tau + \lambda_1(t + (n - 1)\tau) + \lambda_2(t + (n - 1)\tau) \right]
+ \mathcal{M}(2\tau, X)E_{t+(n-1)\tau} \left[ p_M(2\tau, Y) \left[ r^D(t + n\tau)\tau + \lambda_1^*(t + n\tau) + \lambda_2^*(t + n\tau) \right] \right]
= \mathcal{M}(\tau, X) \left[ r^D(t + (n - 1)\tau)\tau + \lambda_1(t + (n - 1)\tau) + \lambda_2(t + (n - 1)\tau) + (\chi - 1)\tau \right].
\]
As a result, the choice of capital (44) of the paper in period \( t + (n - 2)\tau \) is \( K^j_L(t + (n - 1)\tau) \) satisfies
\[
\mathcal{M}(\tau, X)(\chi - 1)\tau = \mathcal{M}(2\tau, X)E_{t+(n-2)\tau} \left[ p_M(2\tau, Y) \left( r^D(t + (n - 1)\tau)\tau + \lambda_1(t + (n - 1)\tau) + \lambda_2(t + (n - 1)\tau) + (\chi - 1)\tau \right) \right].
\]
By (30), (58), and (59) the marginal value of capital in period \( t + (n - 2)\tau \) is
\[
\frac{\partial V}{\partial K^j_L(t + (n - 2)\tau)} = \mathcal{M}(\tau, X) \left[ r^D(t + (n - 2)\tau)\tau + \lambda_1(t + (n - 2)\tau) + \lambda_2(t + (n - 2)\tau) \right]
+ \mathcal{M}(2\tau, X)E_{t+(n-2)\tau} \left[ p_M(2\tau, Y) \left[ r^D(t + (n - 1)\tau)\tau + \lambda_1^*(t + (n - 1)\tau) + \lambda_2^*(t + (n - 1)\tau) \right] \right]
= \mathcal{M}(\tau, X) \left[ r^D(t + (n - 2)\tau)\tau + \lambda_1(t + (n - 2)\tau) + \lambda_2(t + (n - 2)\tau) + (\chi - 1)\tau \right]
\]
This result has the same form as (58).
As a result, the choice of capital in period \( t + (n - 3)\tau \) is \( K^j_L(t + (n - 2)\tau) \)
\[
(\chi - 1)\tau = \mathcal{M}(2\tau, X)E_t \left[ p_M(2\tau, Y) \left( r^D(t + (n - 2)\tau)\tau + \lambda_1(t + (n - 2)\tau) \right)
+ \lambda_2(t + (n - 2)\tau) + (\chi - 1)\tau \right].
\]
Thus, in general the optimal condition for the loan desk’s capital is
\[
(\chi - 1)\tau = \mathcal{M}(2\tau, X)E_t \left[ p_M(2\tau, Y) \left( r^D(t + \tau)\tau + \lambda_1(t + \tau) + \lambda_2(t + \tau) + (\chi - 1)\right) \right].
\]
Theorem 5.2. The bank’s choice of capital for the loan desk’s is optimal when (62) holds for all \( n \) and \( \Delta \kappa < 0 \).

This corresponds to Proposition (5.4) in the paper.

We can use the optimal condition (62) to see how the changes in the level, slope and curvature of the yield curve impacts the optimal capital of the loan desk.

\[
\frac{\partial K^j_L(t + \tau)}{\partial K^j_M(t)} = -\frac{1}{\Delta \kappa} \frac{\partial \mathcal{M}(2\tau, X)}{\partial K^j_M(t)} = -\Delta \mathcal{K}_M \Delta \kappa, (63)
\]

where the partial derivatives are given by (50) and (52).

\[
\frac{\partial K^j_L(t + \tau)}{\partial X} = \frac{(\chi - 1)\tau}{\Delta \kappa} \mathcal{M}(2\tau, X) (\sigma_{\mathcal{M}(\tau)})^{-1} \left( X - \mu_{\mathcal{M}(\tau)} \right) - \frac{\Delta X}{\Delta \kappa}, (64)
\]

where the partial derivatives are given by (50) and (53). We can also determine how the capital of the loan desk changes over several periods.

Theorem 5.3. The impulse response of the optimal loan desk’s capital to the level, slope and curvature of the term structure is determined by (63) and (64).

Proof. If the level, slope or curvature increases at time \( t \), then the expected percentage change in the trading desk’s capital at time \( t + \tau \) is \( -(\sigma_{\mathcal{K}(\tau)})^{-1} \left( X - \mu_{\mathcal{K}(\tau)} \right) \) following (34) in the paper. The change in the loan desk’s capital at time \( t + \tau \) is given by (64). At time \( t + 2\tau \) the trading desk’s expected percentage change in the capital is \( -(\sigma_{\mathcal{K}(2\tau)})^{-1} \left( X - \mu_{\mathcal{K}(2\tau)} \right) \).

In addition the change in the loan desk’s capital is now

\[
\frac{\partial E_t(K^j_L(t + 2\tau))}{\partial X} = -E_t \left( \frac{\Delta \mathcal{K}_M(t + 2\tau)}{\Delta \kappa(t + 2\tau)} \right) \mathcal{K}(2\tau, X) (\sigma_{\mathcal{K}(2\tau)})^{-1} \left( X - \mu_{\mathcal{K}(2\tau)} \right)
\]

\[
+ E_t \left( \frac{\partial K^j_L(t + 2\tau)}{\partial X} \right) e^{-A^P(\tau - t)}. \]

Here, the partial derivatives are given by (63) and (64) at time \( 2\tau \).

At time \( t + k\tau \) for \( k \geq 3 \) the trading desk’s expected percentage change in the capital is
\[- \left( \sigma_K(k\tau) \right)^{-1} \left( X - \mu_K(k\tau) \right), \text{ while the change in the loan desk’s capital is} \]

\[
E_t \left( \frac{\partial K_M^j(t+k\tau)}{\partial X} \right) = - E_t \left( \frac{\Delta K_M(t+k\tau)}{\Delta \kappa(t+k\tau)} \right) K(k\tau, X) \left( \sigma_K(k\tau) \right)^{-1} \left( X - \mu_K(k\tau) \right) + E_t \left( \frac{\partial K_L^j(t+k\tau)}{\partial X(t-k\tau)} \right) e^{-A^p((k-1)\tau-t)}.\]

### 5.1 Optimal \( K_M^j \)

To solve the COO’s problem we use the first order conditions (45) to (47) in the paper. Here, the marginal value of the trading desk’s capital is given by

\[
\frac{\partial V}{\partial K_M^j(t)} = \mathcal{M}(\tau, X) \left[ \mathcal{K}(\tau, X) + \frac{\xi p}{\bar{D}} \left[ \bar{D} - \xi \mathcal{K}_M^j(t) \right] \right. \\
\left. + (1 - \xi \alpha_T) \lambda_2(t) \right] + \mathcal{M}(2\tau, X) E_t \left[ p_M(2\tau, Y) \frac{\partial V}{\partial K_M^j(t + \tau)} \right].
\]

In addition, the marginal value of the loan desk is given by (30) given the optimal choice of capital for the trading desk.

\[
\frac{\partial V}{\partial K_L^j(t)} - EMV(X, K_M^j, K_L^j(\tau, K_M^j, X)) = \mathcal{M}(\tau, X) \left[ r^D(t) \tau + \lambda_1(t) + \lambda_2(t) \right]. \quad (65)
\]

Now substitute the marginal value of capital for the loan desk into the marginal value for the trading desk to yield

\[
\frac{\partial V}{\partial K_M^j(t)} = \frac{\partial V}{\partial K_L^j(t)} \left[ EMV(X, K_M^j, K_L^j(\tau, K_M^j, X)) - \mathcal{M}(2\tau, X) E_t \left[ p_M(2\tau, Y) \frac{\partial V}{\partial K_M^j(t + \tau)} \right] \right] \\
+ \mathcal{M}(\tau, X) \left[ \mathcal{K}(\tau, X) + \frac{\xi p}{\bar{D}} \left[ \bar{D} - \xi \mathcal{K}_M^j(t) \right] \right. \\
\left. - \xi r^D(t) \tau - \xi \kappa T \lambda_1(t) - \xi \alpha_T \lambda_2(t) \right]. \quad (66)
\]

If the bank has both a trade and loan desk in the future, then \( \frac{\partial V}{\partial K_M^j(t+\tau)} = \frac{\partial V}{\partial K_L^j(t+\tau)} = 0 \), then

\[
\frac{\partial V}{\partial K_M^j(t)} - \frac{\partial V}{\partial K_L^j(t)} = \mathcal{M}(\tau, X) \left[ \mathcal{K}(\tau, X) + \frac{\xi p}{\bar{D}} \left[ \bar{D} - \xi \mathcal{K}_M^j(t) \right] \right. \\
\left. - \xi r^D(t) \tau - \xi \kappa T \lambda_1(t) - \xi \alpha_T \lambda_2(t) \right]. \quad (67)
\]
This corresponds to (72) in the paper. If we set this equation to zero using (47) in the paper. (73) in the paper follows from setting this equation to zero using (47) in the paper.

Consider this condition at $K_M^{jN}$ given by (71) in the paper.

$$\frac{\partial V}{\partial K_M^j(t)} - \frac{\partial V}{\partial K_L^j(t)} = \mathcal{M}(\tau, X) \left[ K(\tau, X) + \xi \frac{r_T}{2D} \left[ \bar{D} - \xi K_M^{jN}(t) \right] - \xi r_T^D(t) \tau - \xi \kappa_T \lambda_1(t) - \xi \alpha_T \lambda_2(t) \right] \geq 0$$

$$= [K(\tau, X) + \xi \frac{r_T}{2} - \xi \frac{r_T}{2} \bar{D} \left[ \frac{2}{\xi^2 r_T} K(\tau, X) + 1 \right] - \xi r_T^D(t) \tau - \xi \kappa_T \lambda_1(t) - \xi \alpha_T \lambda_2(t)] \geq 0$$

$$\Rightarrow r_T^D(t) \geq 2 \left[ r_T^D(t) + \kappa_T \lambda_1(t) + \alpha_T \lambda_2(t) \right].$$

(68)

This corresponds to (74) in the paper.

Suppose the capital constraint is binding, then

$$\frac{\partial V}{\partial K_M^j(t)} - \frac{\partial V}{\partial K_L^j(t)} = K(\tau, X) + \xi \frac{r_T}{2D} \left[ \bar{D} - \xi K_M^{jN}(t) \right] - \xi (d_0 + d_1X) - \frac{2 \xi \chi \tau}{\kappa_L (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j)} \left\{ \frac{\gamma_0, \tau + \sigma_0 \varepsilon_{l,t}^j}{2} \right\}$$

$$\Rightarrow K_M^j(t) = \frac{2 \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j) \bar{D}}{\xi^2 [2 \kappa_T \chi \bar{D} \tau + r_T \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j)]} \left( K(\tau, X) + \frac{r_T}{2} - \xi (d_0 + d_1X) \right)$$

$$- \frac{4 \xi \chi \kappa_L \bar{D} \tau}{\xi^2 [2 \kappa_T \chi \bar{D} \tau + r_T \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j)]} \left\{ \frac{\gamma_0, \tau + \sigma_0 \varepsilon_{l,t}^j}{2} - \frac{1}{\kappa_L} \left[ K^j(t) - c_b (P(t, X) - 1) \right] \right\}$$

$$- \frac{1}{2} (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j) \left( c^j + d_0 + d_1X \right) \right\}.$$

This corresponds to (77) in the paper.

We can also find how changes in factors influence the capital for the trading desk.

$$\frac{\partial K_M^j(t)}{\partial X} = \frac{2 \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j) \bar{D}}{\xi^2 [2 \kappa_T \chi \bar{D} \tau + r_T \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j)]} \left( K(\tau, X) (\sigma_K(\tau))^{-1} \left( X - \mu_K(\tau) \right) + \xi d_1 \right)$$

$$- \frac{4 \xi \chi \kappa_L \bar{D} \tau}{\xi^2 [2 \kappa_T \chi \bar{D} \tau + r_T \kappa_L^2 (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j)]} \left\{ \frac{1}{\kappa_L} c_b (P(t, X)b_{sr} | X < X) - \frac{1}{2} (\gamma_{1,\tau}^j - \sigma_1 \varepsilon_{l,t}^j) d_1 \right\}.$$

(69)
If the capital constraint is not binding, then

\[
\frac{\partial V}{\partial K_j^M(t)} - \frac{\partial V}{\partial K_j^L(t)} = K_j^M(\tau,X) + \xi \frac{r}{2D} \left[ D - \xi K_j^M(t) \right] - \xi(d_0 + d_1X) = 0
\]

(70)

\[\Rightarrow K_j^M(t) = \frac{2D}{\xi^2 r} \left( K_j^M(\tau,X) + \frac{\xi r}{2} - \xi(d_0 + d_1X) \right) .\]

The impact of X is

\[
\frac{\partial K_j^M(t)}{\partial X} = -\frac{2D}{\xi^2 r} \left( K_j(\tau,X) (\sigma K(\tau))^{-1} \left( X - \mu K(\tau) \right) + \xi d_1 \right).
\]

(71)

6 Calibrating The Bank Parameters

This section explains how the parameters of the simulation in the paper are determined. First we set the parameters for the liquidity (20) and capital (21) constraints. We then determine the parameters for the bank using the 500 largest commercial banks in the United States from Quarter I 2001 to Quarterly IV 2007. Finally, we provide evidence on the relation between monetary policy and the yield curve factors.

6.1 Regulatory Constraints

To determine the parameters for the regulatory constraint we use Michael King (2010, pp. 10-11) who provides a simplified model of the NSFR.

\[
NSFR = \frac{Equity + Debt_{>1yr} + Liabs_{>1yr} + 0.85StableDeposits_{<1yr} + 0.70OtherDeposits}{0.05GovtDebt + 0.50CorpLoans_{<1yr} + 0.85RetLoans_{<1yr} + OtherAssets} > 1.
\]

(72)

which implies

\[
Equity + Debt_{>1yr} + Liabs_{>1yr} + 0.85StableDeposits_{<1yr} + 0.70OtherDeposits \\
> 0.05GovtDebt + 0.50CorpLoans_{<1yr} + 0.85RetLoans_{<1yr} + OtherAssets.
\]
In our model we have

\[ K^j \equiv \text{Equity} \]

\[ L^j_t = 0.50 \text{CorpLoans}_{<1yr} + 0.85 \text{RetLoans}_{<1yr} \]

\[ L^j_{2\tau,t} + L^j_{2\tau,t-\tau} \equiv \text{Other Assets} \]

\[ D^j \equiv \text{Stable Deposits}_{<1yr} \]

\[ OL^j \equiv \text{Debt}_{>1yr} + \text{Liabs}_{>1yr} + 0.70 \text{Other Deposits} \]

\[ \omega K^j_M \equiv \text{GovtDebt} \]

These variables are stated relative to total assets in Table 9 for the average of the 574 bank holding companies with more than $1 billion in assets as of March 31, 2015. Bank Capital is Total Equity capital. Short term loans, \( L^j_t \) is the Commercial and Industrial Loans plus Loans to Individuals minus Automobile loans. We subtract off Auto Loans since they tend to be longer than one year. For GovtDebt, \( \omega K^j_M \) we use U. S. Government Securities plus Securities Issued by states and Political Subdivisions. For bank reserves, \( R^j_t \) we take Currency and Coin in Domestic Offices plus Balances due from Federal Reserve Banks. For \( \text{Stable Deposits}_{<1yr} \) we use Transaction Accounts + NonTransaction Accounts - Total Time Deposits.

In terms of our model of the NSFR we take\(^5\)

\[ K^j_t \geq 0.05 \omega K^j_M + 0.675 L^j_{\tau,t} + L^j_{2\tau,t} + L^j_{2\tau,t-\tau} - 0.85 D^j - \text{Other Liabilities} . \quad (73) \]

The 0.68 is based on equal weight for corporate and retail loans.

From the balance sheet constraint we have

\[ D^j = R^j + L^j_{\tau,t} + 0.675 L^j_{2\tau,t} + L^j_{2\tau,t-\tau} - \text{Other Liabilities} - K^j_t. \]

We place the same weight on short term loans. Otherwise these loans would reduce capital requirements.

\[ K^j_t \geq 0.027 \left( T^j_{\tau,t} P_{\tau,t} + T^j_{2\tau,t} P_{2\tau,t} + T^j_{3\tau,t} P_{3\tau,t} + T^j_{4\tau,t} P_{4\tau,t} \right) + 0.055 L^j_{\tau,t} + 0.08 \left( L^j_{2\tau,t} + L^j_{2\tau,t-\tau} \right) - 0.459 R^j - 0.069 \text{Other Liabilities} . \quad (74) \]

\(^5\)Since commercial and retail loans are consolidated in the model we take the weight to be the average of 0.5 and 0.85.
$$K^j_t \geq \alpha \tau L^j_t + \alpha_2 \tau (L^j_{2, t} + L^j_{2, t-\tau}) - \alpha_R R^j_t + \alpha_T (T^j_{1, t} P_{1, t} + T^j_{2, t} P_{2, t} + T^j_{3, t} P_{3, t} + T^j_{4, t} P_{4, t}).$$

(75)

The parameters for the regulatory constraints are given in Table 10. This Table corresponds to Table 3 in the paper.

Table 9: Average Accounting Ratios for Commercial Banks with more than a Billion $ (March 31, 2015).

<table>
<thead>
<tr>
<th>Variable</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>$K^j_t$</td>
<td>11.22</td>
</tr>
<tr>
<td>$L^j_t$</td>
<td>15.59</td>
</tr>
<tr>
<td>$L^j_{2,t} + L^j_{2,t-\tau}$</td>
<td>58.82</td>
</tr>
<tr>
<td>$R^j_t$</td>
<td>10.4</td>
</tr>
<tr>
<td>$(T^j_{2,t} P_{2,t} + T^j_{3,t} P_{3,t})$</td>
<td>15.19</td>
</tr>
<tr>
<td>$D^j_t$</td>
<td>56.24</td>
</tr>
<tr>
<td>$O^j_t$</td>
<td>32.54</td>
</tr>
</tbody>
</table>

Table 10: Parameters for Regulatory Constraints (74) and (75).

<table>
<thead>
<tr>
<th>$\alpha_\tau$</th>
<th>$\alpha_2 \tau$</th>
<th>$\alpha_R$</th>
<th>$\alpha_T$</th>
<th>$\kappa_T$</th>
<th>$\kappa_L$</th>
<th>$\bar{b}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.055</td>
<td>0.08</td>
<td>0.459</td>
<td>0.027</td>
<td>0.0</td>
<td>0.08</td>
<td>0.02</td>
</tr>
</tbody>
</table>

6.2 Benchmark Parameters for Banking Model

Next we identify the parameters for the bank. To set the deposit rate parameters in

$$r^{Dj}_{\tau,t} = d_0 + d_1 X_1(t) + \epsilon_{r,Dj,t},$$

we use the data from 500 largest U. S. Commercial Banks from 2001 Quarter I to 2007 Quarter IV. A panel regression of the interest expense on deposits relative to the first interest rate factor $X_1(t)$ is provided in Table 11. Bank fixed effects are included in the regression. We
Table 11: Panel Regression of Deposit Rate on First State Variable.

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>Constant</th>
<th>$X(1)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>0.0111</td>
<td>0.0282</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(7.4196)</td>
<td>(4.6694)</td>
</tr>
<tr>
<td>$adjR^2$</td>
<td>0.2487</td>
<td></td>
</tr>
</tbody>
</table>

Table 12: Panel Regression of Bank Reserves/Assets on First State Variable.

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>Constant</th>
<th>$X_1(t)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>0.1340</td>
<td>0.3936</td>
</tr>
<tr>
<td>T-Stat</td>
<td>17.4619</td>
<td>12.6935</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.5541</td>
<td></td>
</tr>
</tbody>
</table>

Table 13: Parameters for Deposits (51) and Reserves (52) in the Paper.

<table>
<thead>
<tr>
<th>$d_0$</th>
<th>$d_1$</th>
<th>$r_0$</th>
<th>$r_1$</th>
<th>$c_j$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.0111</td>
<td>0.0282</td>
<td>0.1340</td>
<td>0.3936</td>
<td>0.0378</td>
</tr>
</tbody>
</table>

use interest expense on total deposits relative to total deposits, which has a mean value of 0.0078. We use the estimates in Table 11 for the deposit rate parameters in Table 13.

To find parameters for the bank reserves we run the panel regression with bank fixed effects, and using the same set of banks and time period.

$$R^j_t = r_0 + r_1X_1(t) + \epsilon_{R,j,t}.$$  

We estimate this relation using a panel regression with bank fixed effects in Table 12 for the 500 largest U.S. Commercial Banks from 2001 Quarter I to 2007 Quarter IV. The dependent variable is cash balances plus deposits due from other depository institutions including the Federal Reserve. These estimates are included in the Table 13 for the parameters used in the simulations. In the paper this Table corresponds to Table 7.

For reserves we use cash balances plus deposits due from other depository institutions divided by total assets for the 500 largest commercial banks using data from 2001 to 2007. This number includes balances due from depository institutions which is not part of reserves.
Yet, cash items in process of collection plus balances due from Federal Reserve divided by total assets, 10.83%, is 80% of cash balances plus deposits due from other depository institutions divided by total assets for the commercial banks with more than one Billion $ as of March 31, 2015.\(^6\)

We estimate the demand for loans (48) in the paper.

\[
L_{\tau,t}^{d,j} = \gamma_{0,\tau}^{j} - \gamma_{1,\tau}^{j} r_{\tau,t}^{j} + \sigma(r_{\tau,t}^{j}) \varepsilon_{\tau,t}^{j} \text{ with } \sigma(r_{\tau,t}^{j}) = \sigma_{0} + \sigma_{1} r_{\tau,t}^{j}.
\]

To estimate the loan demand we use Commercial and Industrial Loans divided by bank assets for the 500 largest Commercial Banks from 2001 Quarter I to 2007 Quarter IV. Table 14 contains the results for a panel regression with bank fixed effects. We also control for bank size by including the logarithm of bank assets.

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>Constant</th>
<th>(r_{\tau,t}^{j})</th>
<th>(\ln(\text{Assets}))</th>
</tr>
</thead>
<tbody>
<tr>
<td>(\beta)</td>
<td>0.373368</td>
<td>-0.02971</td>
<td>-0.017113</td>
</tr>
<tr>
<td>T-Stat</td>
<td>22.9841</td>
<td>(4.2635)</td>
<td>15.4891</td>
</tr>
<tr>
<td>(\text{adj} R^2)</td>
<td>0.8739</td>
<td>S.E.</td>
<td>0.0331</td>
</tr>
</tbody>
</table>

The slope of the demand curve is \(-\gamma_{1,\tau}^{j} + \sigma_{1} \varepsilon_{\tau,t}^{j}\). As a result, the inverse of the elasticity of demand is

\[
-\frac{1}{\epsilon} = \frac{\frac{\partial r_{\tau,t}^{j}}{\partial L_{\tau,t}^{d,j} \cdot r_{\tau,t}^{j}}}{\frac{1}{\gamma_{0,\tau}^{j} + \sigma_{0} \varepsilon_{\tau,t}^{j} + \left(-\gamma_{1,\tau}^{j} + \sigma_{1} \varepsilon_{\tau,t}^{j}\right) r_{\tau,t}^{j}} + \frac{\gamma_{0,\tau}^{j} + \sigma_{0} \varepsilon_{\tau,t}^{j}}{r_{\tau,t}^{j}(-\gamma_{1,\tau}^{j} + \sigma_{1} \varepsilon_{\tau,t}^{j})} < 0
\]

\[
MR = \left(1 + \frac{\partial r_{\tau,t}^{j}}{\partial L_{\tau,t}^{d,j} \cdot r_{\tau,t}^{j}}\right) r_{\tau,t}^{j} = r_{\tau,t}^{j} \left(1 - \frac{1}{\epsilon}\right) = 2r_{\tau,t}^{j} - \frac{\gamma_{0,\tau}^{j} + \sigma_{0} \varepsilon_{\tau,t}^{j}}{\left(\gamma_{1,\tau}^{j} - \sigma_{1} \varepsilon_{\tau,t}^{j}\right)}.
\]

(76)

Here, \(\frac{1}{\epsilon} = -\frac{\partial r_{\tau,t}^{j}}{\partial L_{\tau,t}^{d,j} \cdot r_{\tau,t}^{j}}\) is the inverse of elasticity of demand. When loans are not constrained

\(^6\)We divided the parameters from the level of reserves regression by the total average across all banks and time periods.
the first order condition is

\[
\frac{\partial \pi_{L}^j}{\partial m_{\tau,t}^j} = \left[ 2 \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}^j \right) r_{\tau,t}^j - \left( c^j + r_{\tau,t}^D \right) \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}^j \right) + \gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t}^j \right] \tau
\]

\[
= \left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}^j \right) \left[ -\left( c^j + r_{\tau,t}^D \right) + 2r_{\tau,t}^j + \frac{\gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t}^j}{\left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}^j \right)} \right] = 0
\]

\[
\Rightarrow r_{\tau,t}^j = \frac{1}{2} \left( c^j + r_{\tau,t}^D \right) + \frac{\gamma_{0,\tau} + \sigma_0 \varepsilon_{\tau,t}^j}{\left( -\gamma_{1,\tau} + \sigma_1 \varepsilon_{\tau,t}^j \right)}.
\]

(77)

We pick the values of the parameters \( \gamma_{0,\tau} \) and \( \gamma_{1,\tau} \) using the average data across the 500 U.S. Commercial Banks from Quarter I of 2001 to Quarter IV of 2007. We want to match the average value of the commercial and industrial loans relative to assets \( L_{\tau}^j/A_{\tau}^j = 0.1212 \), the loan rate of \( r_{\tau,t}^j = 0.0643 \) using the ratio of interest income and fees to loans for commercial and industrial loans. \( c^j = 0.0376 \), for the average non-interest expenses divided by total assets. The average interest expense on deposits to total deposit ratio is used to set \( r_{\tau,t}^D = 0.0165 \). Finally, we want the coefficients to yield the optimal unconstrained loan rate of \( r_{\tau,t}^j = 0.0643 \) when the uncertainty is zero. This leads to the relations:

1. \[ \frac{\gamma_{0,\tau}}{2(\gamma_{1,\tau})} = r_{\tau,t}^j - \frac{1}{2} \left( c^j + r_{\tau,t}^D \right) = 0.03719. \]

2. \[ \gamma_{0,\tau} = \gamma_{1,\tau} r_{\tau,t}^j + L^j = \gamma_{1,\tau} r_{\tau,t}^j + L^j = \frac{\gamma_{0,\tau}}{2(0.03719)} r_{\tau,t}^j + L^j. \]

3. Yields \( \gamma_{0,\tau} = 0.8972 \) and \( \gamma_{1,\tau} = 12.0621 \).

4. The Marginal Revenue is \( 2r_{\tau,t}^j - \frac{\gamma_{0,\tau}}{\gamma_{1,\tau}} = 2 \times 0.0643 - \frac{0.1526}{0.0516} = 0.0543 \).

Table 15: Parameters for Loan Demand (48) and (49) in the paper.

<table>
<thead>
<tr>
<th>( \gamma_{0,\tau} )</th>
<th>( \gamma_{1,\tau} )</th>
<th>( \sigma_0 )</th>
<th>( \sigma_1 )</th>
<th>( z_0 )</th>
<th>( z_1 )</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.8972</td>
<td>12.0621</td>
<td>0.0331</td>
<td>0.2067</td>
<td>-0.6150</td>
<td>0.00035</td>
</tr>
</tbody>
</table>

If we use the parameters from the regression we have \( \gamma_0 = 0.1121 \) and \( \gamma_1 = 0.030 \). However, Marginal revenue turns out to be too big, since \( \gamma_0/\gamma_1 = 3.7367 \). Consequently, we choose the parameters in Table 15 which are based on the unconstrained profit maximization condition and the loan demand curve.
From the 500 Largest U.S. Commercial banks we have total charge offs to total assets is 0.00566 with a standard deviation of 0.0133. The standard error of the regressions in Table 14 is 0.0331. We want the standard error $\sigma_0 + \sigma_1 r^j_{\tau,t} = 0.0331 + 0.0133 = 0.0464$ so that $\sigma_0 = 0.0331$ and $\sigma_1 = 0.0133/r^j_{\tau,t} = 0.2067$ Consider a two point distribution with payoff $z_0$ with probability $p = 0.00566$ and $z_1$ with probability $1 - p$. Suppose we know mean $x$ and variance $y$, then basic algebra shows

$$z_0 = x - \sqrt{\frac{1-p}{p}} y$$

$$z_1 = x + \sqrt{\frac{p}{1-p}} y.$$ 

These leads to the solutions $z_0 = -0.6150$ and $z_1 = 0.0035$. $r^j_{\tau,t}(z_0) = 0.0631$ and $r^j_{\tau,t}(z_1) = 0.0643$ when the interest rate is kept at its optimum. This completes Table 15 which is reproduced in Table 8 of the paper.

### 6.3 Yields Factors and Macroeconomic Variables

Recent work by Joslin, Priebsch and Singleton (2014) has examined the relation between the principle components of yields data and macroeconomic variables. They find that the level is positively affected by economic growth and inflation. This corresponds to the usual result that interest rates increase during booms to the business cycle and when inflation increases. At the same time a higher level of economic growth leads to a flatter slope for a positive sloped yield curve, since a central bank would want to raise short term interest rates when the economy is expanding too fast. In addition, they find that higher inflation increases the slope of the yield curve by a smaller magnitude relative to economic growth. While the third principle component is not affected by economic growth and inflation, it does have a negative impact on economic growth and inflation. Consequently, higher curvature signals lower economic growth and inflation, which in turn leads to lower level and a larger slope for the yield curve. Thus, there is a significant connection among the latent factors and economic growth and inflation, which corresponds to how the yield curve behaves over the business cycle.

We examine such connections between yields factors and economic variables using our data in the sample period 1990M01-2013M12. Specifically, we base our analysis within the well established monetary policy framework of the Taylor rule (1993). Since we illustrate the implications of the model using a one factor model, we investigate how a one factor model

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\[^7\text{They use the Chicago Federal Reserve index of economic activity for economic growth.}\]
captures monetary policy. To this end, we first regress the short term 3 Month Treasury yield on the deviation of inflation $\Delta P_{t-1}$ from its target $\Delta P^*$, as well as the GDP gap $Y^{gap} = \frac{Y_P - Y}{Y_P}$.

\[
r_{3mo,t} = \Delta P_{t-1} + r_t^* + a_{\Delta P}(\Delta P_{t-1} - \Delta P^*) + a_Y Y^{gap}.
\]

To measure the Taylor variables we use Real Potential Gross Domestic Product, Billions of Chained 2009 Dollars, Quarterly, Not Seasonally Adjusted from FRED, $Y^P$ and Monthly real GDP from Macroeconomic Analysis http://www.macroadvisers.com/monthly-gdp/, $Y$. We interpolate the quarterly potential GDP to obtain the monthly observations. For inflation, $\Delta P$ we use Consumer Price Index for All Urban Consumers: All Items Less Food and Energy, Change from Year Ago, Index 1982-84=100, Monthly, Seasonally Adjusted from FRED. Finally, we use the 3-month yield to maturity on Treasury security as a proxy for the short term rate.\(^8\) As a result, $Y^{gap}$ is negative during a recession. $r_t^*$ is the natural real rate of interest. In addition, the Taylor principle assumes that $a_{\Delta P} = 0.5$ and $a_Y = 0.5$. As a result a 1\% increase in inflation leads to a $1\frac{1}{2}$\% increase in the short term interest rate, while a 1\% increase in the output gap leads to a $\frac{1}{2}$\% increase in short term interest rates.

The results of the Taylor rule regression are in Table 16. Note that for this exercise and all of the following exercises using factors, the t-test statistics are calculated using Newey-West HAC (Heteroskedasticity and Autocorrelation Consistent) estimator because the yields data are typically highly persistent. The regression is run for data from April 1992 to December 2013 because monthly real GDP starts from April 1992. We obtain the expected signs for the impact of inflation and the GDP gap. Yet, the Taylor principle does not hold since a change in the inflation rate leads to a less than 1\% increase in short term interest rates, so that the Federal Reserve is less aggressive in combating inflation relative to the Taylor principle. Furthermore, the response of the short rate to inflation rate is not significant.\(^9\)

Because the short rate variation reflects both the level and slope movements of the yield curve, we study how each of the factors responds to the macroeconomic variables using both the empirically constructed factors and the term structure latent factors obtained through Kalman filter within the affine term structure model. Both constructions are presented in the previous subsection.

\(^8\)We could also use the Federal Funds rate and obtain similar results.

\(^9\)Note, however, that we use lagged variables in the Taylor-rule regression here as a simple specification to illustrate how yields factors are generally related to the economic variables.
Table 16: OLS Regression of 3 Month yield on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>$Y_{t-1}^{gap}$</th>
<th>$\Delta P_{t-1}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>0.6005</td>
<td>0.4299</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(7.1273)</td>
<td>(1.1759)</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.6641</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.0993</td>
<td></td>
</tr>
</tbody>
</table>

First we focus on the empirical factors as constructed in the previous section by following Diebold and Li (2006). Recall the level explains 95.5% of the variation of the yield curve, while slope and curvature explain 4.2% and 0.2% of this variation, respectively. We estimate the Taylor rule regressions using the level, slope and curvature factors of the yield curve as the dependent variable and the results are given in Tables 17, 18 and 19. First, the level factor regression yields quantitatively similar results to that using the 3-month short rate in Table 16. A positive and significant response coefficient estimate associated with the output gap suggests that the level factor tends to rise (decline) in response to economic booms (recessions). Although the response coefficient estimate associated with the inflation rate is also consistent with economic intuitions it is again not significant and the magnitude does not satisfy the Taylor principle. The adjusted R-square of the level factor regression turns out to be higher than those of both the slope and curvature factors regressions. Despite a lower adjusted R-square, the slope factor regression also yields results consistent with economic intuitions. The negative and significant response coefficient estimate of the output gap suggests that the Fed tends to cut (raise) the short rate leading to larger (smaller) slope factor during recessions (booms). Again, although the response coefficient estimate of the inflation rate is consistent with Taylor rule it is not significant. Turning to the curvature factor, the regression results reveal that it appears to be significantly positively correlated with the output gap.

For comparison purposes we also run the same Taylor rule regressions using the latent factors obtained in the affine term structure model via the Kalman Filter, and the results are given in Tables 20, 21, and 22. The Taylor rule regressions results are broadly similar to those using the empirical factors. However, the adjusted R-square for the first latent factor that is closely related to the level factor becomes lower than those for the second and third factors.

Overall, the above regression results suggest yield curve factors contain useful information
Table 17: OLS Regression of Level on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>$Y_{t-1}^{gap}$</th>
<th>$\Delta P_{t-1}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>0.4648</td>
<td>0.2615</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(4.1191)</td>
<td>(0.9649)</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.4991</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.0814</td>
<td></td>
</tr>
</tbody>
</table>

Table 18: OLS Regression of Slope on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>$Y_{t-1}^{gap}$</th>
<th>$\Delta P_{t-1}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>-0.1357</td>
<td>-0.1684</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(-2.8873)</td>
<td>(-0.6095)</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.2134</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.1381</td>
<td></td>
</tr>
</tbody>
</table>

Table 19: OLS Regression of Curvature on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>$Y_{t-1}^{gap}$</th>
<th>$\Delta P_{t-1}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\beta$</td>
<td>0.1438</td>
<td>0.1307</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(4.2511)</td>
<td>(0.7840)</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.3337</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.1417</td>
<td></td>
</tr>
</tbody>
</table>

about how monetary policy reacts to economic conditions. For example, during economic booms, consistently the level factor rises and the slope factor declines. This suggests that during economic booms when the Fed raises the policy rate, it shifts up the whole yield curve leading to a rise of the level factor, and at the same time, such a policy raises the short end of the yield curve more than the long end so that the yield curve also becomes flatter. These observations suggest that it is entirely possible to use one factor such as the level factor to capture the monetary policy.
Table 20: OLS Regression of First Latent Variable on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>( Y_{t-1}^{gap} )</th>
<th>( \Delta P_{t-1} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \beta )</td>
<td>0.1657</td>
<td>0.0899</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(1.2073)</td>
<td>(0.2284)</td>
</tr>
<tr>
<td>( R^2 )</td>
<td>0.0967</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.0504</td>
<td></td>
</tr>
</tbody>
</table>

Table 21: OLS Regression of Second Latent Variable on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>( Y_{t-1}^{gap} )</th>
<th>( \Delta P_{t-1} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \beta )</td>
<td>-0.5046</td>
<td>-0.1869</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(-5.2797)</td>
<td>(-1.0252)</td>
</tr>
<tr>
<td>( R^2 )</td>
<td>0.5000</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.2640</td>
<td></td>
</tr>
</tbody>
</table>

Table 22: OLS Regression of Third Latent Variable on Taylor Rule Variables

<table>
<thead>
<tr>
<th>Variable/Statistic</th>
<th>( Y_{t-1}^{gap} )</th>
<th>( \Delta P_{t-1} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \beta )</td>
<td>0.9083</td>
<td>0.5300</td>
</tr>
<tr>
<td>T-Stat</td>
<td>(8.2229)</td>
<td>(1.0401)</td>
</tr>
<tr>
<td>( R^2 )</td>
<td>0.6308</td>
<td></td>
</tr>
<tr>
<td>D-W</td>
<td>0.1591</td>
<td></td>
</tr>
</tbody>
</table>